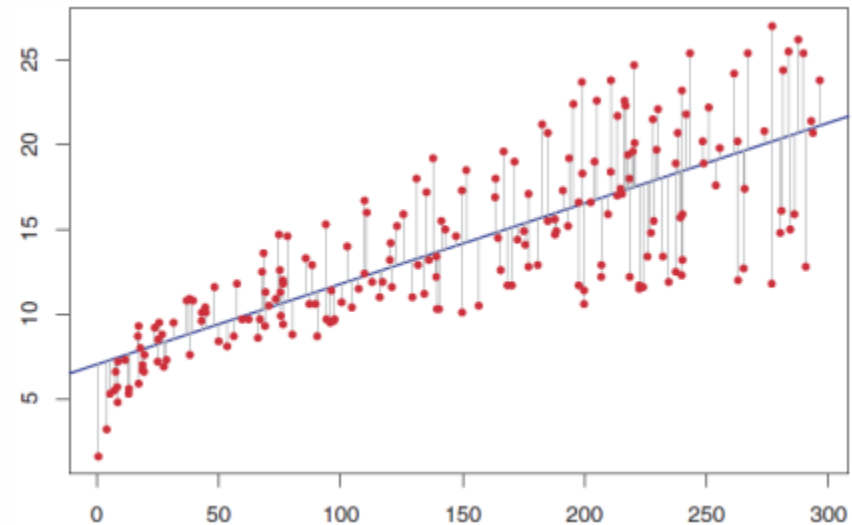




Pattern Analysis and Machine Intelligence

Linear Regression

- Simple Linear Regression Model
 - Least Squares Fit
 - Measures of Fit
 - Inference in Regression
- Multi Variate Regression Model
 - Least Squares Fit
 - Inference in Regression
- Other Considerations in Regression Model
 - Qualitative Predictors
 - Interaction Terms
 - Non Linear Regression
 - Possible issues in Linear Regression



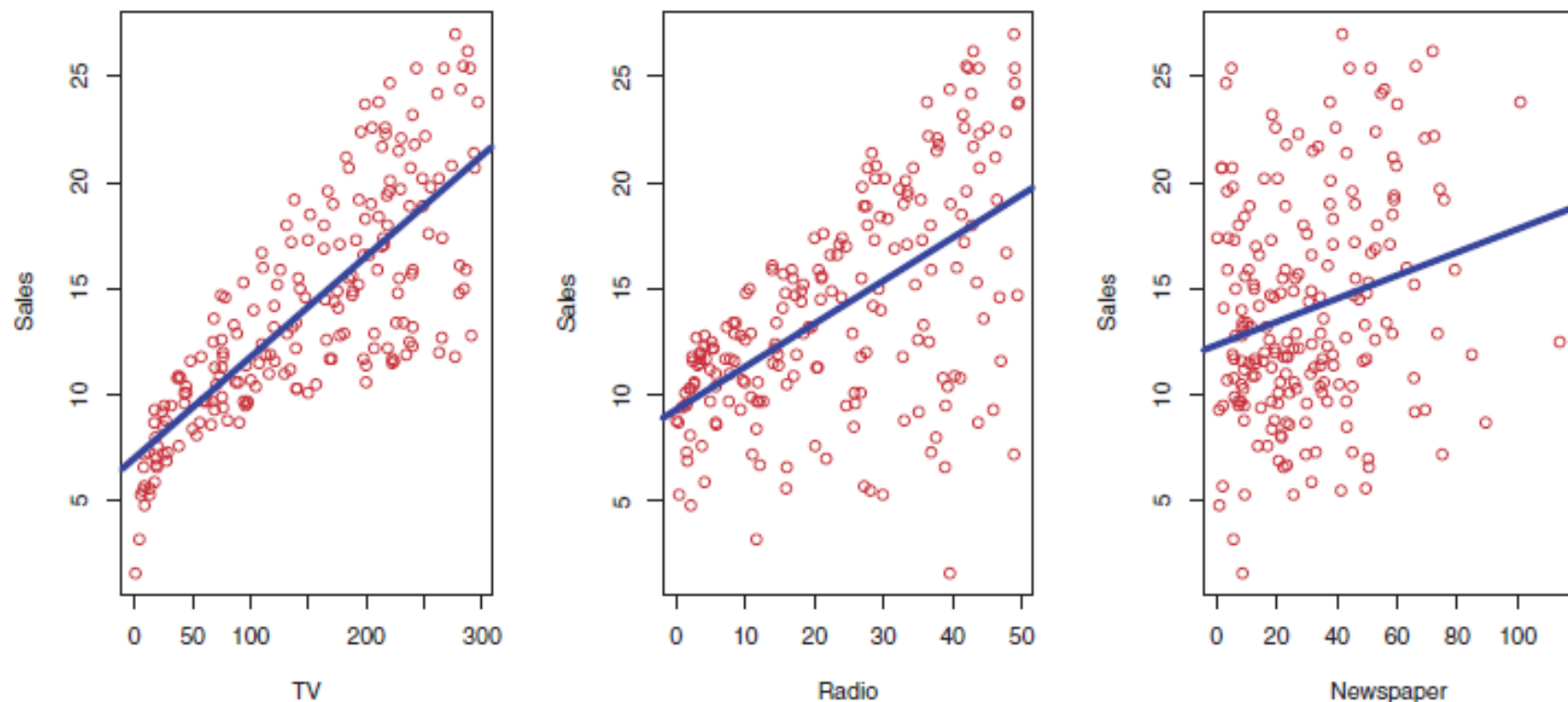
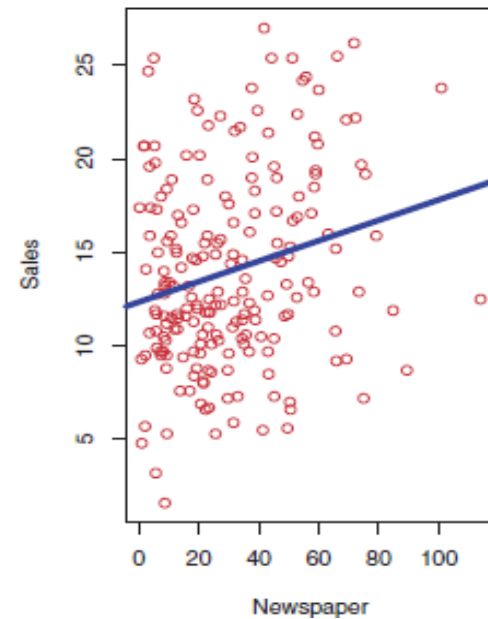
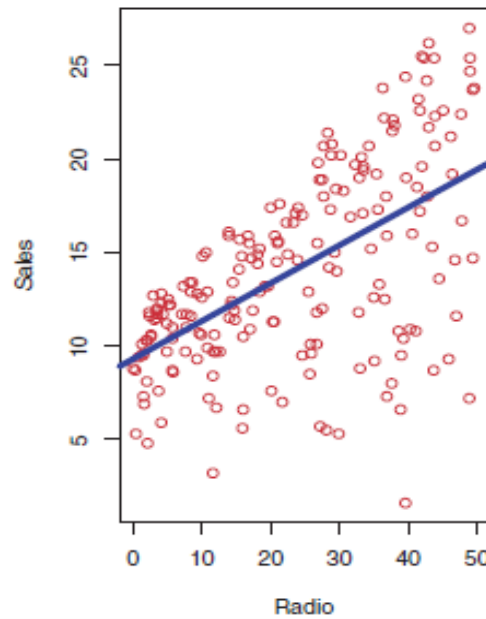
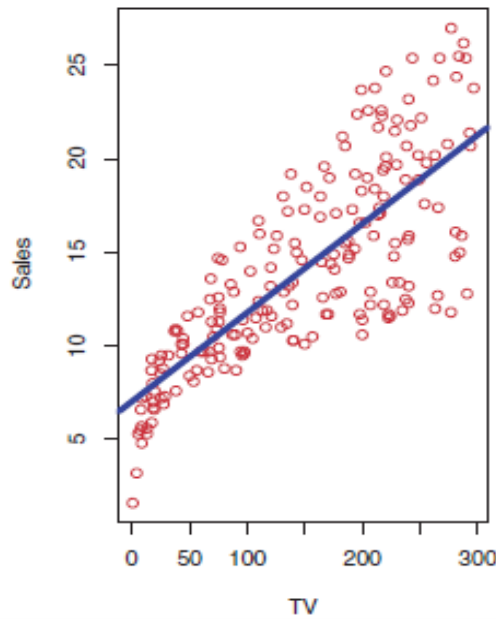


FIGURE 2.1. The Advertising data set. The plot displays sales, in thousands of units, as a function of TV, radio, and newspaper budgets, in thousands of dollars, for 200 different markets. In each plot we show the simple least squares fit of sales to that variable, as described in Chapter 3. In other words, each blue line represents a simple model that can be used to predict sales using TV, radio, and newspaper, respectively.

What can we ask to the data?

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- *Is there a relationship between advertising budget and sales?*
- *How strong is the relationship between advertising budget and sales?*
- *Which media contribute to sales?*
- *How accurately can we estimate the effect of each medium on sales?*
- *How accurately can we predict future sales?*
- *Is the relationship linear?*
- *Is there synergy among the advertising media?*

- Let's assume that a linear relationship exists between Y and X

$$\text{sales} \approx \beta_0 + \beta_1 \times \text{TV}$$

- We say sales regress on TV through some parameters
 - Model coefficients β_0 and β_1
 - After training, a new data point can be predicted as

$$\hat{y} = \hat{\beta}_0 + \hat{\beta}_1 x$$

- Given a dataset the coefficient above can be estimated by using least squares to minimize the Residual Sum of Squares

$$e_i = y_i - \hat{y}_i$$

$$\text{RSS} = e_1^2 + e_2^2 + \dots + e_n^2$$

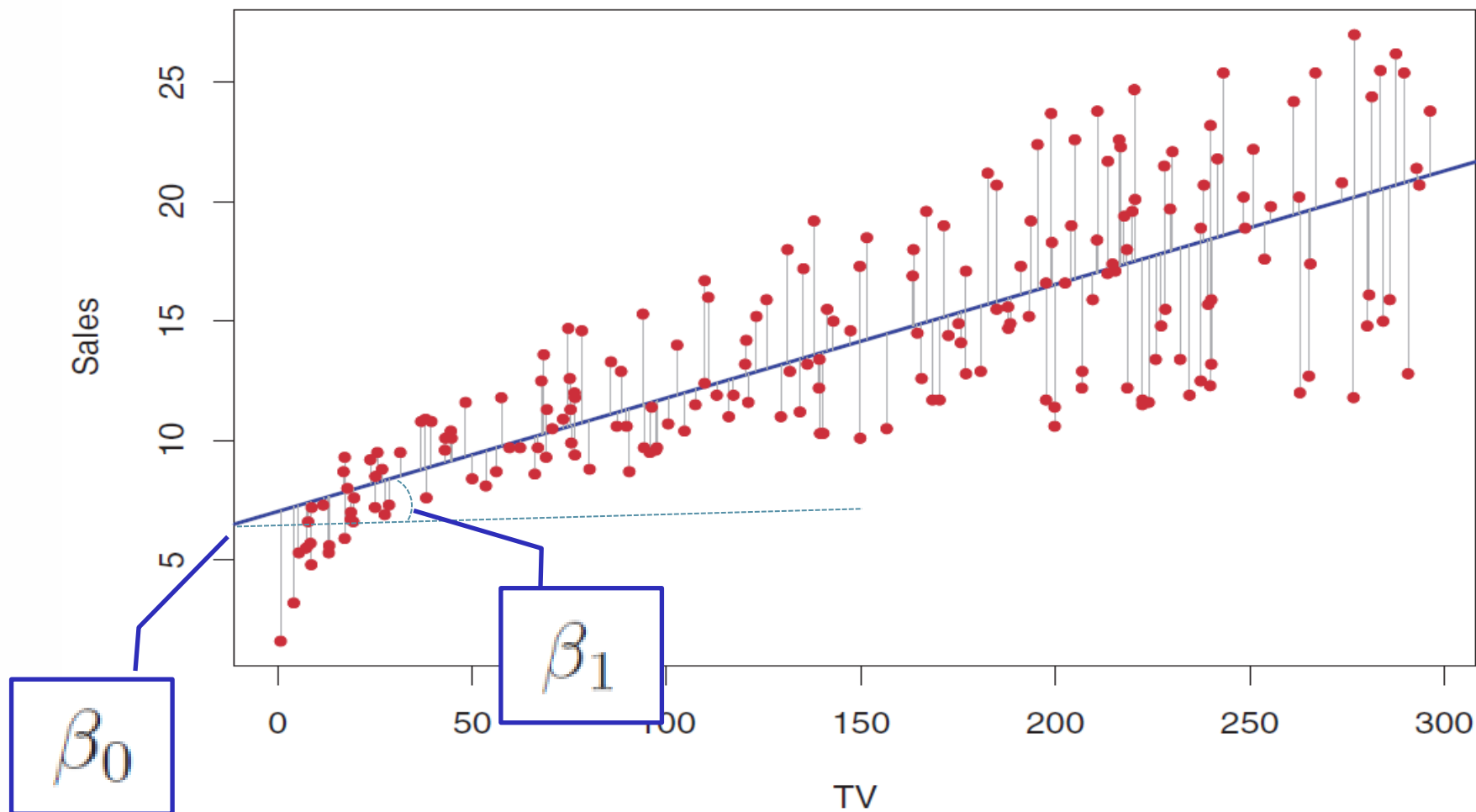


FIGURE 3.1. For the **Advertising** data, the least squares fit for the regression of **sales** onto **TV** is shown. The fit is found by minimizing the sum of squared errors. Each grey line segment represents an error, and the fit makes a compromise by averaging their squares. In this case a linear fit captures the essence of the relationship, although it is somewhat deficient in the left of the plot.

- Least square fitting minimizes RSS (Residual Sum of Squares)

$$e_i = y_i - \hat{y}_i$$

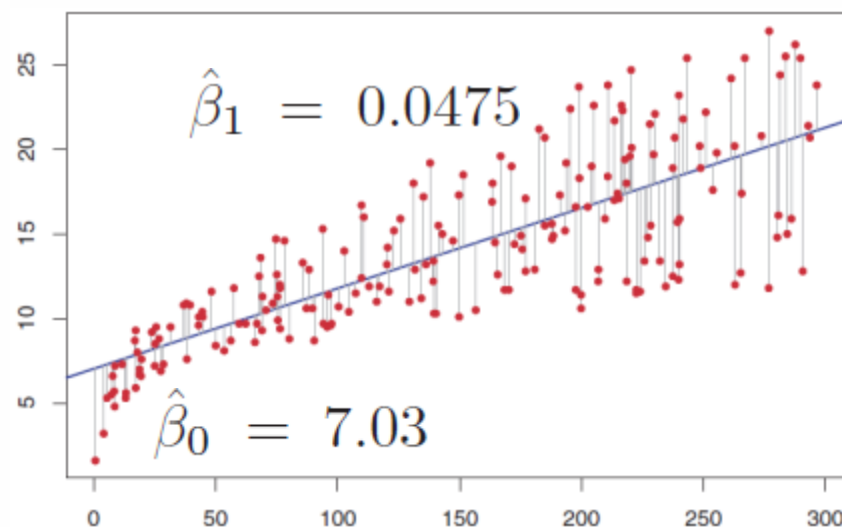
$$\text{RSS} = e_1^2 + e_2^2 + \dots + e_n^2$$

$$\text{RSS} = (y_1 - \hat{\beta}_0 - \hat{\beta}_1 x_1)^2 + (y_2 - \hat{\beta}_0 - \hat{\beta}_1 x_2)^2 + \dots + (y_n - \hat{\beta}_0 - \hat{\beta}_1 x_n)^2$$

- Obtaining the following estimates

$$\hat{\beta}_1 = \frac{\sum_{i=1}^n (x_i - \bar{x})(y_i - \bar{y})}{\sum_{i=1}^n (x_i - \bar{x})^2}$$

$$\hat{\beta}_0 = \bar{y} - \hat{\beta}_1 \bar{x}$$



- Where $\bar{y} \equiv \frac{1}{n} \sum_{i=1}^n y_i$ and $\bar{x} \equiv \frac{1}{n} \sum_{i=1}^n x_i$

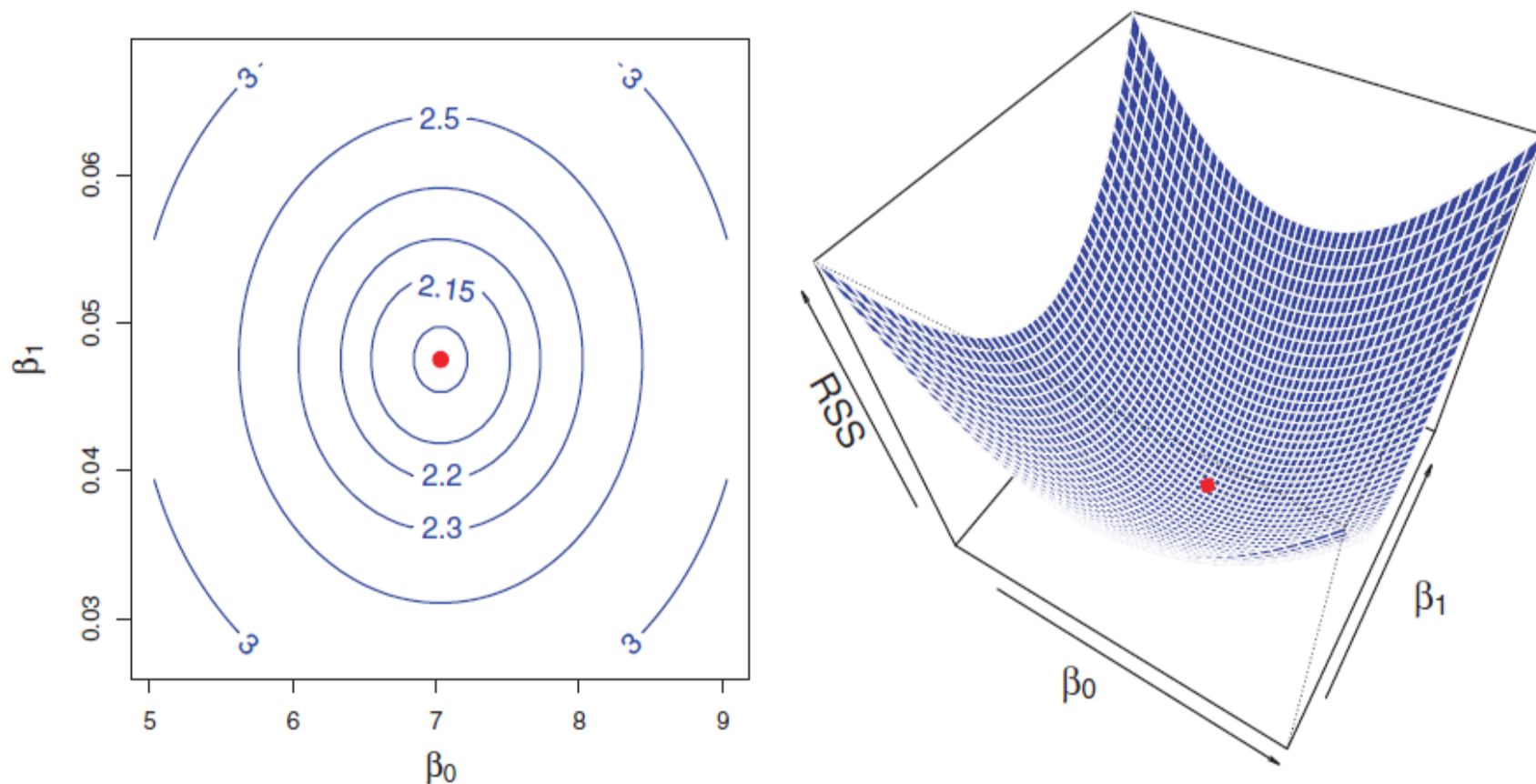


FIGURE 3.2. Contour and three-dimensional plots of the RSS on the Advertising data, using sales as the response and TV as the predictor. The red dots correspond to the least squares estimates $\hat{\beta}_0$ and $\hat{\beta}_1$, given by (3.4).

- Least square fitting minimizes RSS (Residual Sum of Squares)

$$e_i = y_i - \hat{y}_i$$

$$\text{RSS} = e_1^2 + e_2^2 + \dots + e_n^2$$

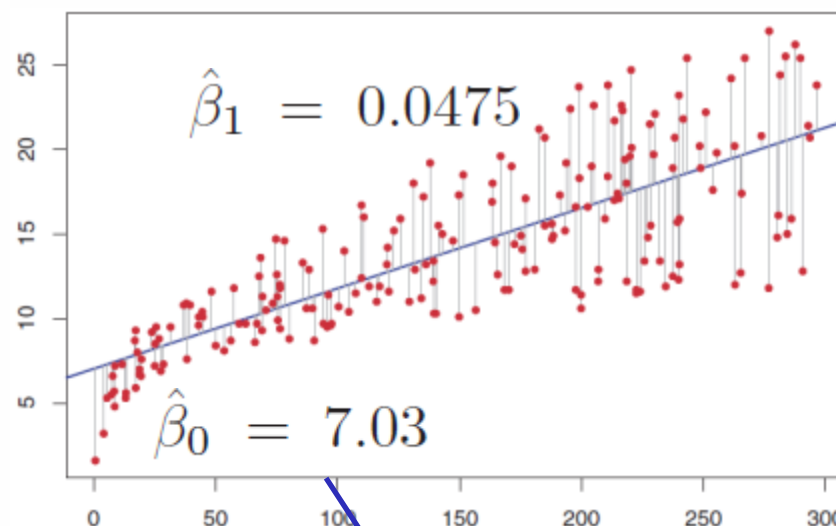
$$\text{RSS} = (y_1 - \hat{\beta}_0 - \hat{\beta}_1 x_1)^2 + (y_2 - \hat{\beta}_0 - \hat{\beta}_1 x_2)^2 + \dots + (y_n - \hat{\beta}_0 - \hat{\beta}_1 x_n)^2$$

- Obtaining the following estimates

$$\hat{\beta}_1 = \frac{\sum_{i=1}^n (x_i - \bar{x})(y_i - \bar{y})}{\sum_{i=1}^n (x_i - \bar{x})^2}$$

$$\hat{\beta}_0 = \bar{y} - \hat{\beta}_1 \bar{x}$$

- Where $\bar{y} \equiv \frac{1}{n} \sum_{i=1}^n y_i$ and $\bar{x} \equiv \frac{1}{n} \sum_{i=1}^n x_i$



*But are they
any good?*

- Recall from Statistical Learning theory the underlying hypothesis

$$Y = \beta_0 + \beta_1 X + \epsilon.$$

- β_0 the Y value when $X = 0$
 - β_1 the average increase in Y due to unitary increase in X
 - the error term captures all the rest ...
- This model is known as “population regression line”
 - the best linear approximation of the true model
 - it might differ from the least squares regression line
- “The population regression line stays to the mean of a distribution as the least squares regression line stays to the sample mean ...”*

Example: population regression line

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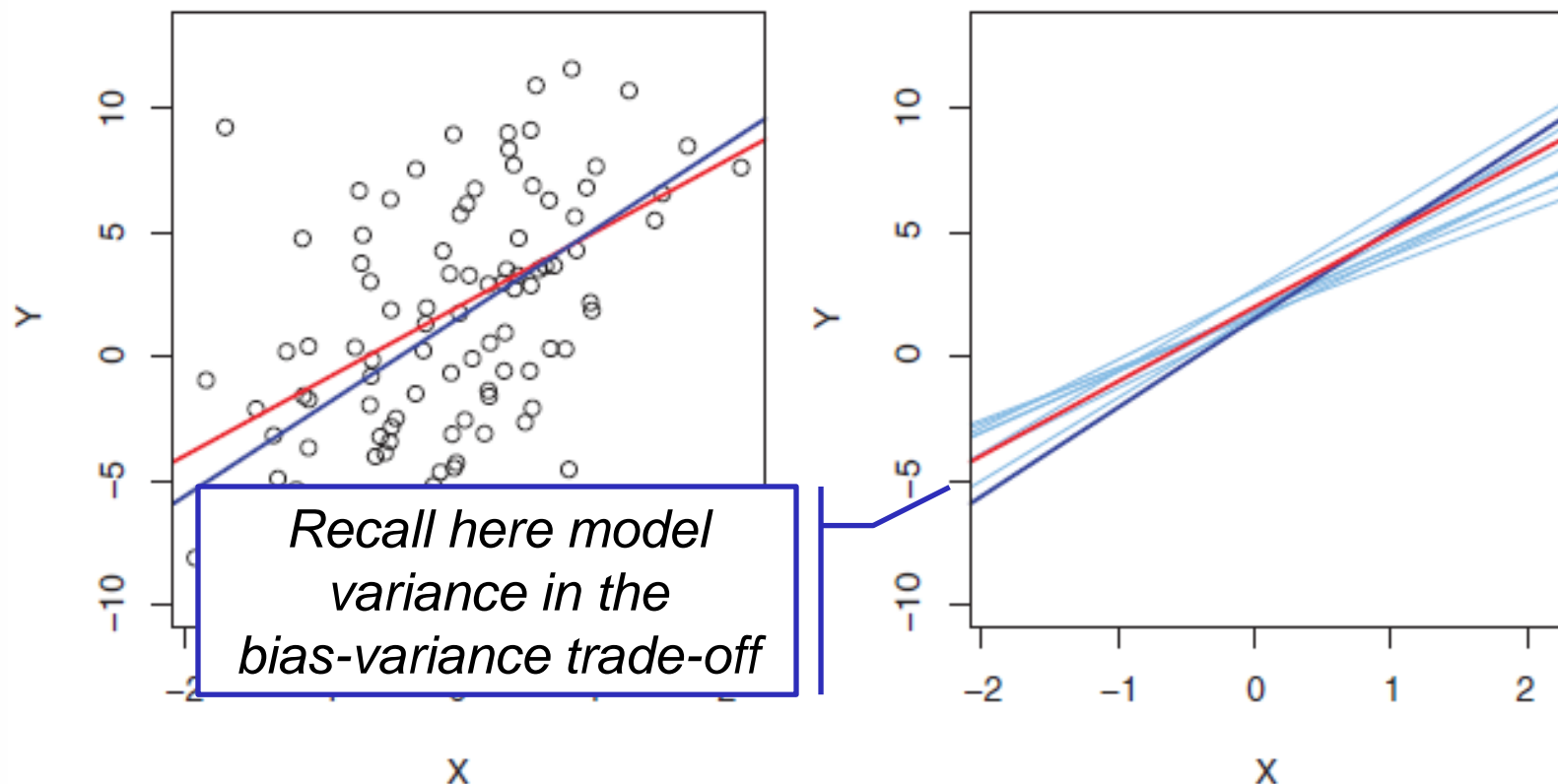


FIGURE 3.3. A simulated data set. Left: The red line represents the true relationship, $f(X) = 2 + 3X$, which is known as the population regression line. The blue line is the least squares line; it is the least squares estimate for $f(X)$ based on the observed data, shown in black. Right: The population regression line is again shown in red, and the least squares line in dark blue. In light blue, ten least squares lines are shown, each computed on the basis of a separate random set of observations. Each least squares line is different, but on average, the least squares lines are quite close to the population regression line.

- The (squared) standard error for the mean estimator represents the average distance of the sample mean from the real mean

$$\text{Var}(\hat{\mu}) = \text{SE}(\hat{\mu})^2 = \frac{\sigma^2}{n}$$

- We can use similar formulae for the standard errors of the linear regression coefficients ...

$$\text{SE}(\hat{\beta}_0)^2 = \sigma^2 \left[\frac{1}{n} + \frac{\bar{x}^2}{\sum_{i=1}^n (x_i - \bar{x})^2} \right]$$

$$\text{SE}(\hat{\beta}_1)^2 = \frac{\sigma^2}{\sum_{i=1}^n (x_i - \bar{x})^2}$$

- These formulae assume
 - uncorrelated errors ...

*The higher the spread of x
the better the estimate*

- ... having the same (unknown) variance $\sigma^2 = \text{Var}(\epsilon)$

- In general errors variance is not known, but it can be estimated from residuals (if the model fits properly)

$$\text{RSE} = \sqrt{\text{RSS}/(n - 2)}$$

- From standard errors we can compute confidence intervals for the linear regression parameters.

- E.g., the 95% confidence intervals for the parameters are

$$\hat{\beta}_0 \pm 2 \cdot \text{SE}(\hat{\beta}_0)$$

$$\hat{\beta}_1 \pm 2 \cdot \text{SE}(\hat{\beta}_1)$$

This should be the 97.5 quantile of a t-distribution

- the true slope is, with 95% probability, in the range

$$\left[\hat{\beta}_1 - 2 \cdot \text{SE}(\hat{\beta}_1), \hat{\beta}_1 + 2 \cdot \text{SE}(\hat{\beta}_1) \right]$$

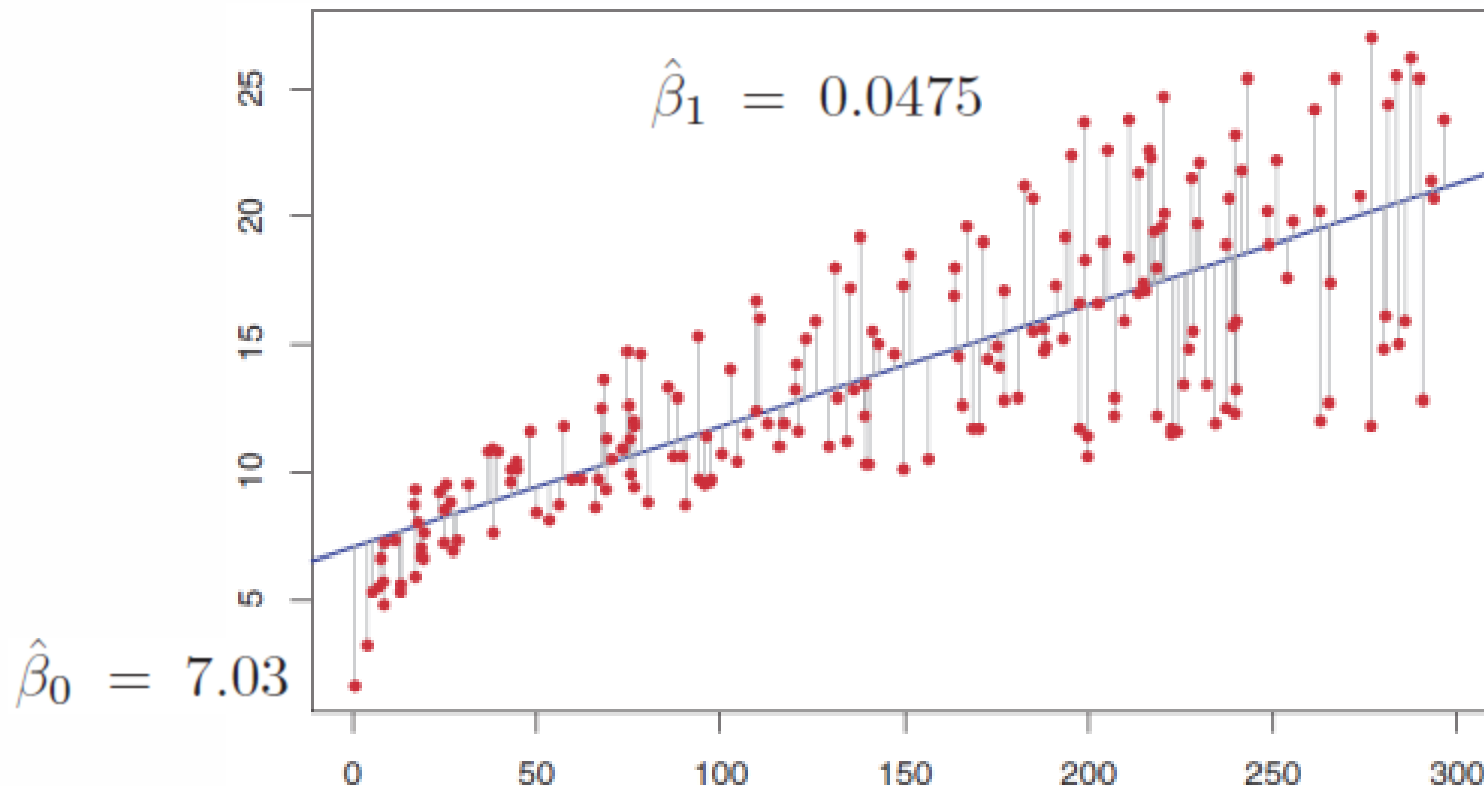
Example: TV Advertising data

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- If we consider the 95% confidence intervals
 - for the intercept we have $[6.130, 7.935]$
 - for the slope we have $[0.042, 0.053]$

Sales without any advertising

Average impact of TV advertising



- Standard errors can be used for hypothesis testing such as:
 - H_0 : there is no relationship between Y and X
 - H_a : there is some relationship between Y and X

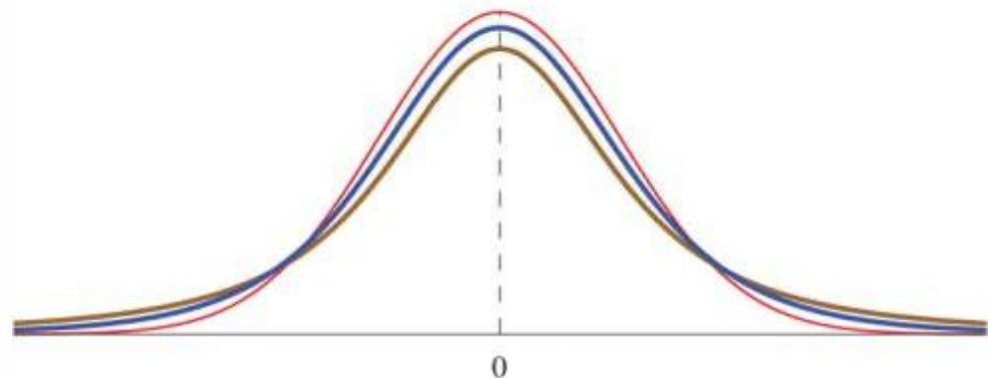
- This translates on parameters hypothesis testing for

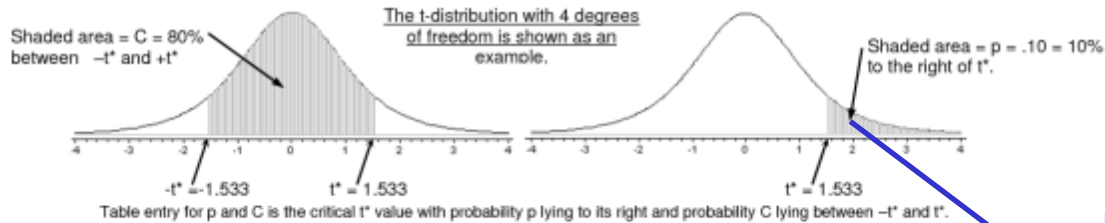
$$H_0 : \beta_1 = 0 \text{ against } H_a : \beta_1 \neq 0$$

- We do not know true parameters so we can use estimates and perform a statistical test using

$$t = \frac{\hat{\beta}_1 - 0}{\text{SE}(\hat{\beta}_1)}$$

t -distribution with $df = 5$
 t -distribution with $df = 2$





Upper Tail Probability $p \rightarrow$	0.25	0.20	0.15	0.10	0.05	0.025	0.02	0.01	0.005	0.0025	0.001	0.0005
Degrees of freedom \downarrow												
1	1.000	1.378	1.963	3.078	6.314	12.71	15.89	31.82	63.66	127.3	318.3	636.6
2	0.818	1.061	1.386	1.886	2.920	4.303	4.849	8.965	9.925	14.09	22.33	31.60
3	0.765	0.978	1.250	1.638	2.353	3.182	3.482	4.541	5.841	7.453	10.21	12.92
4	0.741	0.941	1.190	1.533	2.132	2.776	2.999	3.747	4.604	5.598	7.173	8.610
5	0.727	0.920	1.156	1.476	2.015	2.571	2.757	3.365	4.032	4.773	5.894	6.869
6	0.718	0.906	1.134	1.440	1.943	2.447	2.612	3.143	3.707	4.317	5.208	5.959
7	0.711	0.896	1.119	1.415	1.895	2.365	2.517	2.998	3.499	4.029	4.785	5.408
8	0.706	0.889	1.108	1.397	1.860	2.306	2.449	2.896	3.355	3.833	4.501	5.041
9	0.703	0.883	1.100	1.383	1.833	2.262	2.398	2.821	3.250	3.690	4.297	4.781
10	0.700	0.879	1.093	1.372	1.812	2.228	2.359	2.784	3.169	3.581	4.144	4.587
11	0.697	0.876	1.088	1.363	1.796	2.201	2.328	2.718	3.106	3.497	4.025	4.437
12	0.695	0.873	1.083	1.356	1.782	2.179	2.303	2.681	3.055	3.426	3.930	4.318
13	0.694	0.870	1.079	1.350	1.771	2.160	2.282	2.650	3.012	3.372	3.852	4.221
14	0.692	0.868	1.076	1.345	1.761	2.145	2.264	2.624	2.977	3.326	3.787	4.140
15	0.691	0.866	1.074	1.341	1.753	2.131	2.249	2.602	2.947	3.286	3.733	4.073
16	0.690	0.865	1.071	1.337	1.746	2.120	2.235	2.583	2.921	3.252	3.686	4.015
17	0.689	0.863	1.069	1.333	1.740	2.110	2.224	2.567	2.898	3.222	3.646	3.965
18	0.688	0.862	1.067	1.330	1.734	2.101	2.214	2.552	2.878	3.197	3.610	3.922
19	0.688	0.861	1.066	1.328	1.729	2.093	2.205	2.539	2.861	3.174	3.579	3.883
20	0.687	0.860	1.064	1.325	1.725	2.086	2.197	2.528	2.845	3.153	3.552	3.850
21	0.686	0.859	1.063	1.323	1.721	2.080	2.189	2.518	2.831	3.135	3.527	3.819
22	0.686	0.858	1.061	1.321	1.717	2.074	2.183	2.508	2.819	3.119	3.505	3.792
23	0.685	0.858	1.060	1.319	1.714	2.069	2.177	2.500	2.807	3.104	3.485	3.768
24	0.685	0.857	1.059	1.318	1.711	2.064	2.172	2.492	2.797	3.091	3.467	3.745
25	0.684	0.856	1.058	1.316	1.708	2.060	2.167	2.485	2.787	3.078	3.450	3.725
26	0.684	0.856	1.058	1.315	1.706	2.056	2.162	2.479	2.779	3.067	3.435	3.707
27	0.684	0.855	1.057	1.314	1.703	2.052	2.158	2.473	2.771	3.057	3.421	3.689
28	0.683	0.855	1.056	1.313	1.701	2.048	2.154	2.467	2.763	3.047	3.408	3.674
29	0.683	0.854	1.055	1.311	1.699	2.045	2.150	2.462	2.756	3.038	3.396	3.660
30	0.683	0.854	1.055	1.310	1.697	2.042	2.147	2.457	2.750	3.030	3.385	3.646
40	0.681	0.851	1.050	1.303	1.684	2.021	2.123	2.423	2.704	2.971	3.307	3.551
50	0.679	0.849	1.047	1.299	1.676	2.009	2.109	2.403	2.678	2.937	3.261	3.496
60	0.679	0.848	1.045	1.296	1.671	2.000	2.099	2.396	2.660	2.915	3.232	3.460
80	0.678	0.846	1.043	1.292	1.664	1.990	2.088	2.374	2.639	2.887	3.195	3.416
100	0.677	0.845	1.042	1.290	1.660	1.984	2.081	2.364	2.626	2.871	3.174	3.390
1000	0.675	0.842	1.037	1.282	1.646	1.962	2.056	2.330	2.581	2.813	3.058	3.306
z^*	0.674	0.841	1.036	1.282	1.645	1.960	2.054	2.326	2.576	2.807	3.091	3.291
Confidence level $C = 1 - 2p \rightarrow$	50%	60%	70%	80%	90%	95%	96%	98%	99%	99.5%	99.8%	99.9%

This value is known as *p-value*

$$t = \frac{\hat{\beta}_1 - 0}{SE(\hat{\beta}_1)}$$

$$\hat{\beta}_1 \pm 2 \cdot SE(\hat{\beta}_1)$$

~ 1.984

$N = 200$ data points
 $p\text{-value} = 2.5\%$
 $t \sim 1.984$

Here you find the confidence interval

- We reject the null hypothesis H_0 if the p-value is small
 - p-value is the probability of making a wrong choice
 - usually small is as low as 5% or 1%, these percentages, with $N > 30$ correspond to $t \sim 2$ and $t \sim 2.75$ respectively
 - in other fields, p-values might be significantly different, e.g., in bioinformatics p-values of 10^{-6} are quite common to avoid false discoveries ...

	Coefficient	Std. error	t-statistic	p-value
Intercept	7.0325	0.4578	15.36	< 0.0001
TV	0.0475	0.0027	17.67	< 0.0001

TABLE 3.1. For the **Advertising** data, coefficients of the least squares model for the regression of number of units sold on TV advertising budget. An increase of \$1,000 in the TV advertising budget is associated with an increase in sales by around 50 units (Recall that the **sales** variable is in thousands of units, and the **TV** variable is in thousands of dollars).

- The classical measure of fit is mean squared error, in linear regression we use the Residual Standard Error

$$RSS = \sum_{i=1}^n (y_i - \hat{y}_i)^2$$

$$RSE = \sqrt{\frac{1}{n-2} RSS} = \sqrt{\frac{1}{n-2} \sum_{i=1}^n (y_i - \hat{y}_i)^2}$$

- It estimates the standard deviation of the errors, i.e., the irreducible error.

Quantity	Value
Residual standard error	3.26
R^2	0.612
F-statistic	312.1

How far the model is from least square line on average

*Compared to the average sales
3,260/14,000 = 23 %*

TABLE 3.2. For the **Advertising** data, more information about the model for the regression of number of units sold on TV advertising

- We might be interested in computing how much of the data variance is explained by the model (“relative accuracy”)

$$R^2 = \frac{\text{TSS} - \text{RSS}}{\text{TSS}} = 1 - \frac{\text{RSS}}{\text{TSS}}$$

$$\text{RSS} = \sum_{i=1}^n (y_i - \hat{y}_i)^2$$

$$\text{TSS} = \sum (y_i - \bar{y})^2$$

- An R^2 close to 1 means the data are almost perfectly explained by our simple linear model, in our case it is just 0.612 ...

Quantity	Value
Residual standard error	3.26
R^2	0.612
F-statistic	312.1

Is this due to the error noise or to the fact that data is not linear?

TABLE 3.2. For the **Advertising** data, more information about the least squares model for the regression of number of units sold on TV advertising budget.

- Recall the definition of correlation between two variables, it is a measure of the (linear) relationship between them

$$\text{Cor}(X, Y) = \frac{\sum_{i=1}^n (x_i - \bar{x})(y_i - \bar{y})}{\sqrt{\sum_{i=1}^n (x_i - \bar{x})^2} \sqrt{\sum_{i=1}^n (y_i - \bar{y})^2}}$$

- In the simple univariate linear regression setting we have

$$r = \text{Cor}(X, Y)$$

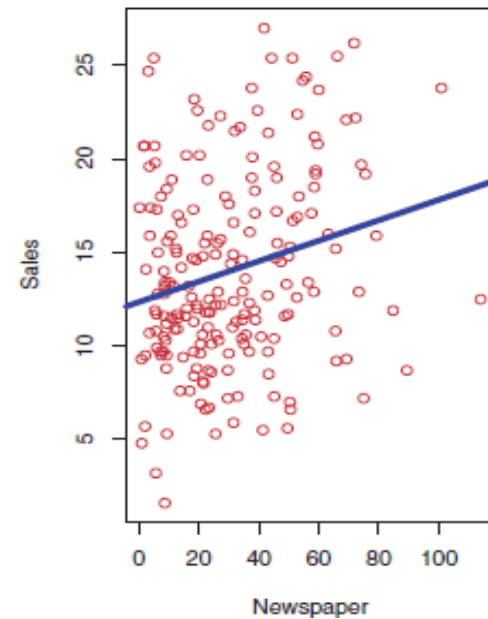
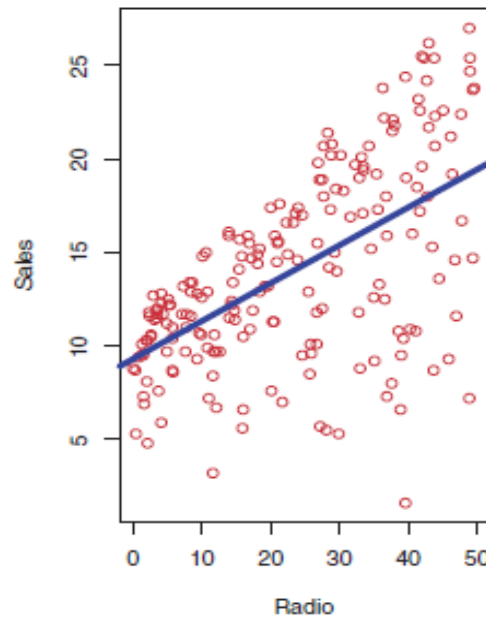
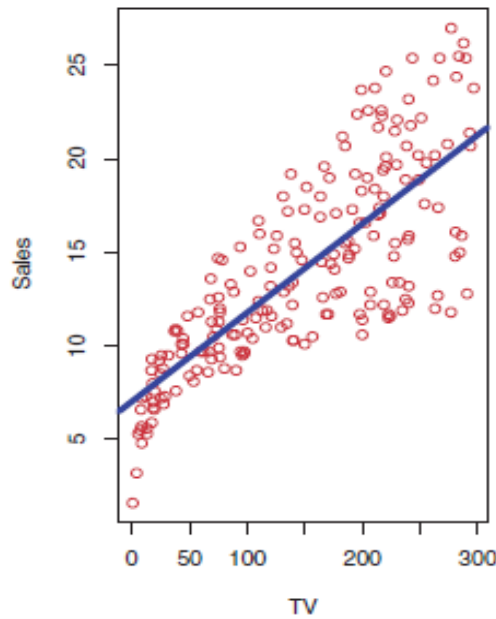
$$R^2 = r^2$$

- In a multivariate case this does not hold and R^2 is used to extend the correlation concept to multiple variables

$$R^2 = \frac{\text{TSS} - \text{RSS}}{\text{TSS}} = 1 - \frac{\text{RSS}}{\text{TSS}}$$

What can we ask to the data?

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- *Is there a relationship between advertising budget and sales?*
- *How strong is the relationship between advertising budget and sales?*
- *Which media contribute to sales?*
- *How accurately can we estimate the effect of each medium on sales?*
- *How accurately can we predict future sales?*
- *Is the relationship linear?*
- *Is there synergy among the advertising media?*

Simple regression of **sales** on **radio**

	Coefficient	Std. error	t-statistic	p-value
Intercept	9.312	0.563	16.54	< 0.0001
radio	0.203	0.020	9.92	< 0.0001

Simple regression of **sales** on **newspaper**

	Coefficient	Std. error	t-statistic	p-value
Intercept	12.351	0.621	19.88	< 0.0001
newspaper	0.055	0.017	3.30	< 0.0001

TABLE 3.3. *More simple linear regression models for the **Advertising** data. Coefficients of the simple linear regression model for number of units sold on Top: radio advertising budget and Bottom: newspaper advertising budget. A \$1,000 increase in spending on radio advertising is associated with an average increase in sales by around 203 units, while the same increase in spending on newspaper advertising is associated with an average increase in sales by around 55 units (Note that the **sales** variable is in thousands of units, and the **radio** and **newspaper** variables are in thousands of dollars).*

- Treating variables as they were independent
 - does not tell how an increase in sales is obtained by changing the three input variables
 - the coefficients of each input did not take into account the other input in the estimate process
 - if input are highly correlated, using independent estimates can be misleading
- Extend linear regression to consider multiple predictors
$$\text{sales} = \beta_0 + \beta_1 \times \text{TV} + \beta_2 \times \text{radio} + \beta_3 \times \text{newspaper} + \epsilon$$
- More formally we have the multivariate regression

$$Y = \beta_0 + \beta_1 X_1 + \beta_2 X_2 + \cdots + \beta_p X_p + \epsilon$$

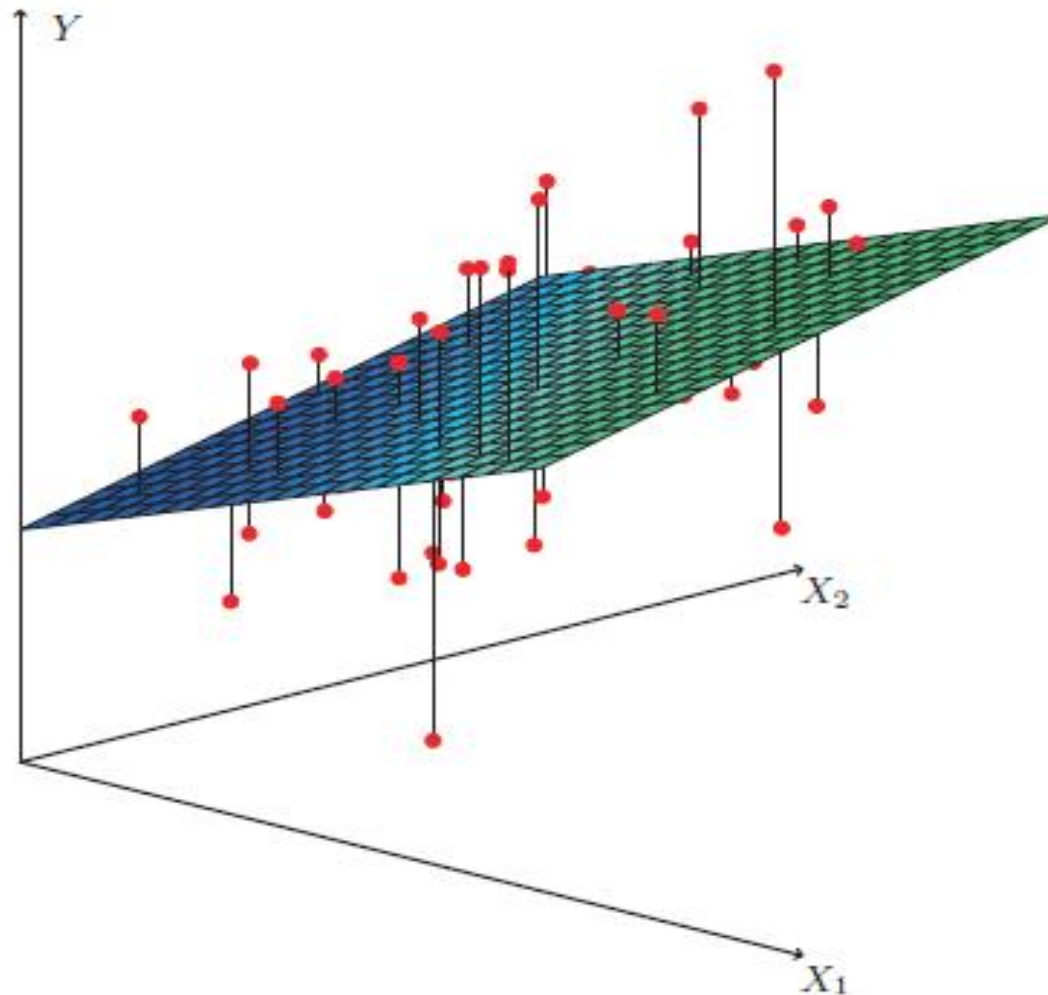


FIGURE 3.4. *In a three-dimensional setting, with two predictors and one response, the least squares regression line becomes a plane. The plane is chosen to minimize the sum of the squared vertical distances between each observation (shown in red) and the plane.*

- Linear regression parametric model, i.e., the population line

$$Y = \beta_0 + \beta_1 X_1 + \beta_2 X_2 + \cdots + \beta_p X_p + \epsilon$$

- each parameters describes the average influence of the associated input keeping all the others fixed
- The regression coefficient can be estimated by least squares fit

$$\begin{aligned} \text{RSS} &= \sum_{i=1}^n (y_i - \hat{y}_i)^2 \\ &= \sum_{i=1}^n (y_i - \hat{\beta}_0 - \hat{\beta}_1 x_{i1} - \hat{\beta}_2 x_{i2} - \cdots - \hat{\beta}_p x_{ip})^2 \end{aligned}$$

- To obtain the least squares predictor

$$\hat{y} = \hat{\beta}_0 + \hat{\beta}_1 x_1 + \hat{\beta}_2 x_2 + \cdots + \hat{\beta}_p x_p$$

	Coefficient	Std. error	t-statistic	p-value
Intercept	7.0325	0.4578	15.36	< 0.0001
TV	0.0475	0.0027	17.67	< 0.0001

	Coefficient	Std. error	t-statistic	p-value
Intercept	2.939	0.3119	9.42	< 0.0001
TV	0.046	0.0014	32.81	< 0.0001
radio	0.189	0.0086	21.89	< 0.0001
newspaper	−0.001	0.0059	−0.18	0.8599

TABLE 3.4. For the Advertising data, least squares coefficient estimates of the multiple linear regression of number of units sold on radio, TV, and newspaper advertising budgets.

	Coefficient	Std. error	t-statistic	p-value
Intercept	9.312	0.563	16.54	< 0.0001
radio	0.203	0.020	9.92	< 0.0001

	Coefficient	Std. error	t-statistic	p-value
Intercept	2.939	0.3119	9.42	< 0.0001
TV	0.046	0.0014	32.81	< 0.0001
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TABLE 3.4. For the **Advertising** data, least squares coefficient estimates of the multiple linear regression of number of units sold on radio, TV, and newspaper advertising budgets.

	Coefficient	Std. error	t-statistic	p-value
Intercept	12.351	0.621	19.88	< 0.0001
newspaper	0.055	0.017	3.30	< 0.0001

	Coefficient	Std. error	t-statistic	p-value
Intercept	2.939	0.3119	9.42	< 0.0001
TV	0.046	0.0014	32.81	< 0.0001
radio	0.189	0.0086	21.89	< 0.0001
newspaper	−0.001	0.0059	−0.18	0.8599

TABLE 3.4. For the Advertising data, least squares coefficient estimates of the multiple linear regression of number of units sold on radio, TV, and newspaper advertising budgets.

	TV	radio	newspaper	sales
TV	1.0000	0.0548	0.0567	0.7822
radio		1.0000	0.3541	0.5762
newspaper			1.0000	0.2283
sales				1.0000

TABLE 3.5. *Correlation matrix for TV, radio, newspaper, and sales for the Advertising data.*

- Let consider correlations between input and output variables
 - If we increase *radio* then *sales* increase
 - *Radio* and *newspaper* are highly correlated in some markets
 - If we increase *radio* then *newspaper* increases
- The increase on sales is correlated to the increase of newspaper is due to radio, not to the fact that newspaper increases sales
 - E.g., increase in shark attacks are correlated to ice cream sales at the beach ... because of people!

- Computing the least regression fit can be done easily using linear algebra calculus

- Recall here
$$\begin{aligned} \text{RSS}(\beta) &= \sum_{i=1}^N (y_i - f(x_i))^2 \\ &= \sum_{i=1}^N \left(y_i - \beta_0 - \sum_{j=1}^p x_{ij} \beta_j \right)^2 \end{aligned}$$

- By taking into account that
 - \mathbf{X} is an $N \times (p+1)$ data matrix
 - \mathbf{y} is $N \times 1$ vector of desired output
 - β is a $(p+1) \times 1$ vector of model coefficients
- We can rewrite the Residuals Sums of Squares as

$$\text{RSS}(\beta) = (\mathbf{y} - \mathbf{X}\beta)^T (\mathbf{y} - \mathbf{X}\beta)$$

- We want to minimize $\text{RSS}(\beta) = (\mathbf{y} - \mathbf{X}\beta)^T (\mathbf{y} - \mathbf{X}\beta)$
- Let's compute the RSS derivatives with respect to

$$\frac{\partial \text{RSS}}{\partial \beta} = -2\mathbf{X}^T (\mathbf{y} - \mathbf{X}\beta)$$

$$\frac{\partial^2 \text{RSS}}{\partial \beta \partial \beta^T} = 2\mathbf{X}^T \mathbf{X}$$

- Assuming \mathbf{X} has full rank and $\mathbf{X}^T \mathbf{X} > 0$ we have just to compare the first derivative to zero

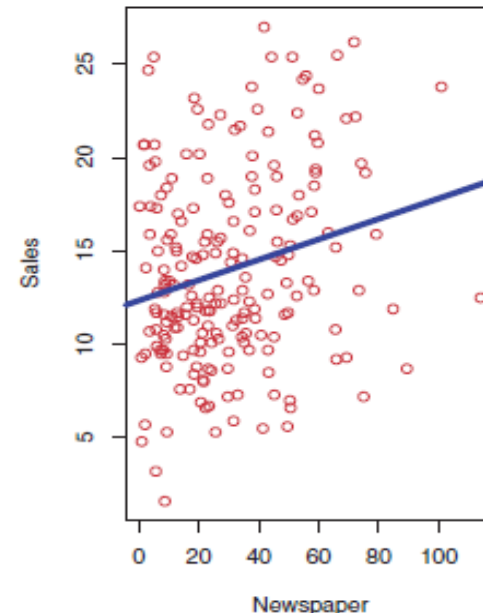
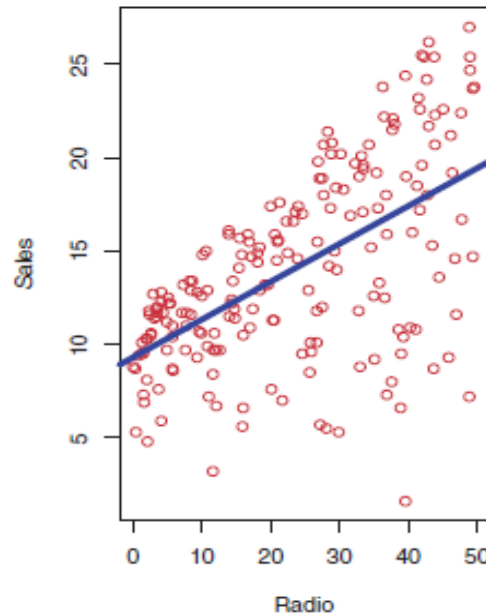
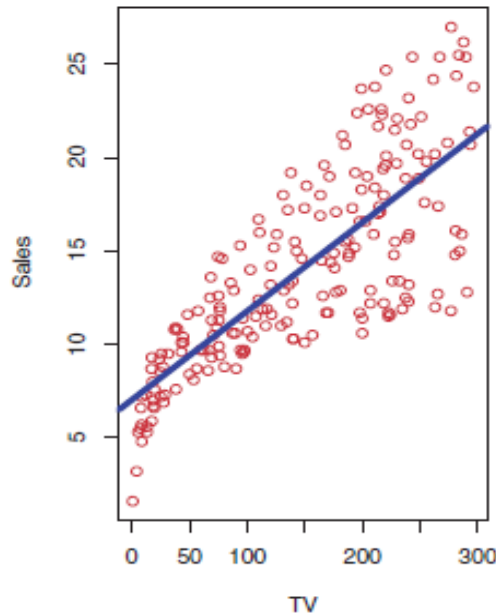
$$\text{Var}(\hat{\beta}) = (\mathbf{X}^T \mathbf{X})^{-1} \sigma^2$$

$$\mathbf{X}^T (\mathbf{y} - \mathbf{X}\beta) = 0$$
$$\hat{\beta} = (\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T \mathbf{y}$$

Pseudo
Inverse

- In matrix algebra terms the prediction becomes

$$\hat{\mathbf{y}} = \mathbf{X}\hat{\beta} = \mathbf{X}(\mathbf{X}^T \mathbf{X})^{-1} \mathbf{X}^T \mathbf{y}$$



- *Is at least one of the predictors X_1, \dots, X_p useful in predicting the response?*
- *Do all the predictors help to explain Y , or it is only a subset of the predictors useful*
- *How well does the model fit the data?*
- *Given a set of predictor values, what response value should we predict, and accurate is our prediction?*

- Is there any relationship between response and predictors?

$$H_0 : \beta_1 = \beta_2 = \dots = \beta_p = 0$$

against

$$H_a : \text{at least one } \beta_j \text{ is non-zero}$$

- This test is performed using the F-statistics

$$F = \frac{(\text{TSS} - \text{RSS})/p}{\text{RSS}/(n - p - 1)}$$

$\text{TSS} = \sum (y_i - \bar{y})^2$

$\text{RSS} = \sum_{i=1}^n (y_i - \hat{y}_i)^2$

- If the linear model assumptions are valid

$$E\{\text{RSS}/(n - p - 1)\} = \sigma^2$$

- when H_0 is true $E\{(\text{TSS} - \text{RSS})/p\} = \sigma^2$ $F \sim I$
- when H_a is true $E\{(\text{TSS} - \text{RSS})/p\} > \sigma^2$ then $F > I$

Quantity	Value
Residual standard error	1.69
R^2	0.897
F-statistic	570

$F > 1$

TABLE 3.6. More information about the least squares model for the regression of number of units sold on TV, newspaper, and radio advertising budgets in the Advertising data. Other information about this model was displayed in Table 3.4.

- F is well above 1, a relationship exists!

F Table for $\alpha = .05$



- How much should be F to tell this relationship exists?

	df ₁ =1	2	3	4	5	6	7	8	9	10
df ₂ =1	161.45	199.50	215.71	224.58	230.16	233.99	236.77	238.88	240.54	241.88
2	18.51	19.00	19.16	19.25	19.30	19.33	19.35	19.37	19.38	19.40
3	10.13	9.55	9.28	9.12	9.01	8.94	8.89	8.85	8.81	8.79
4	7.71	6.94	6.59	6.39	6.26	6.16	6.09	6.04	6.00	5.96
5	6.61	5.79	5.41	5.19	5.05	4.95	4.88	4.82	4.77	4.74
6	5.99	5.14	4.76	4.53	4.39	4.28	4.21	4.15	4.10	4.06
7	5.59	4.74	4.35	4.12	3.97	3.87	3.79	3.73	3.68	3.64
8	5.32	4.46	4.07	3.84	3.69	3.58	3.50	3.44	3.39	3.35
9	5.12	4.26	3.86	3.63	3.48	3.37	3.29	3.23	3.18	3.14
10	4.96	4.10	3.71	3.48	3.33	3.22	3.14	3.07	3.02	2.98
11	4.84	3.98	3.59	3.36	3.20	3.09	3.01	2.95	2.90	2.85
12	4.75	3.89	3.49	3.26	3.11	3.00	2.91	2.85	2.80	2.75
13	4.67	3.81	3.41	3.18	3.03	2.92	2.83	2.77	2.71	2.67
14	4.60	3.74	3.34	3.11	2.96	2.85	2.76	2.70	2.65	2.60
15	4.54	3.68	3.29	3.06	2.90	2.79	2.71	2.64	2.59	2.54
16	4.49	3.63	3.24	3.01	2.85	2.74	2.66	2.59	2.54	2.49
17	4.45	3.59	3.20	2.96	2.81	2.70	2.61	2.55	2.49	2.45

- We can test also a subset of the variables

$$H_0 : \beta_{p-q+1} = \beta_{p-q+2} = \dots = \beta_p = 0$$

- The novel F-statistics for the model fitted on q variables is

$$F = \frac{(\text{RSS}_0 - \text{RSS})/q}{\text{RSS}/(n - p - 1)}$$

- If we leave out one variable at the time ($q=1$) we obtain an equivalent formulation of the t-statistics for single parameters
 - F-statistics is more accurate than t-statistics computed for each parameter since it corrects for other parameters
 - It tells you the partial effect of adding that specific variable to the model

- If the number of factors p is big, p -values might be tricky
 - With $p=100$ and H_0 true, $\sim 5\%$ of the p -values (by chance) will be lower than 0.05 and we might see 5 predictors associated (by chance) to the response
 - F statistic is not affected by the number of factors p in the model

	Coefficient	Std. error	t-statistic	p-value
Intercept	2.939	0.3119	9.42	< 0.0001
TV	0.046	0.0014	32.81	< 0.0001
radio	0.189	0.0086	21.89	< 0.0001
newspaper	-0.001	0.0059	-0.18	0.8599

TABLE 3.4. For the Advertising data, least squares coefficient estimates of the multiple linear regression of number of units sold on radio, TV, and newspaper advertising budgets.

- As for the simple linear regression R^2 and RSE are used
- In multiple linear regression R^2 equals $\text{Cor}(Y, \hat{Y})^2$
- An R^2 close to 1 means that the model explains a large amount of variance in the data
- The more variables in the model the

Removing newspaper this goes from 0.8972 to 0.89719

Quantity	Value
Residual standard error	1.69
R^2	0.897
F-statistic	570

TABLE 3.6. *More information about the least squares model for the regression of number of units sold on TV, newspaper, and radio advertising budgets in the Advertising data. Other information about this model was displayed in Table 3.4.*

- In a multi variable scenario RSE becomes
$$RSE = \sqrt{\frac{1}{n - p - 1} \text{RSS}}$$

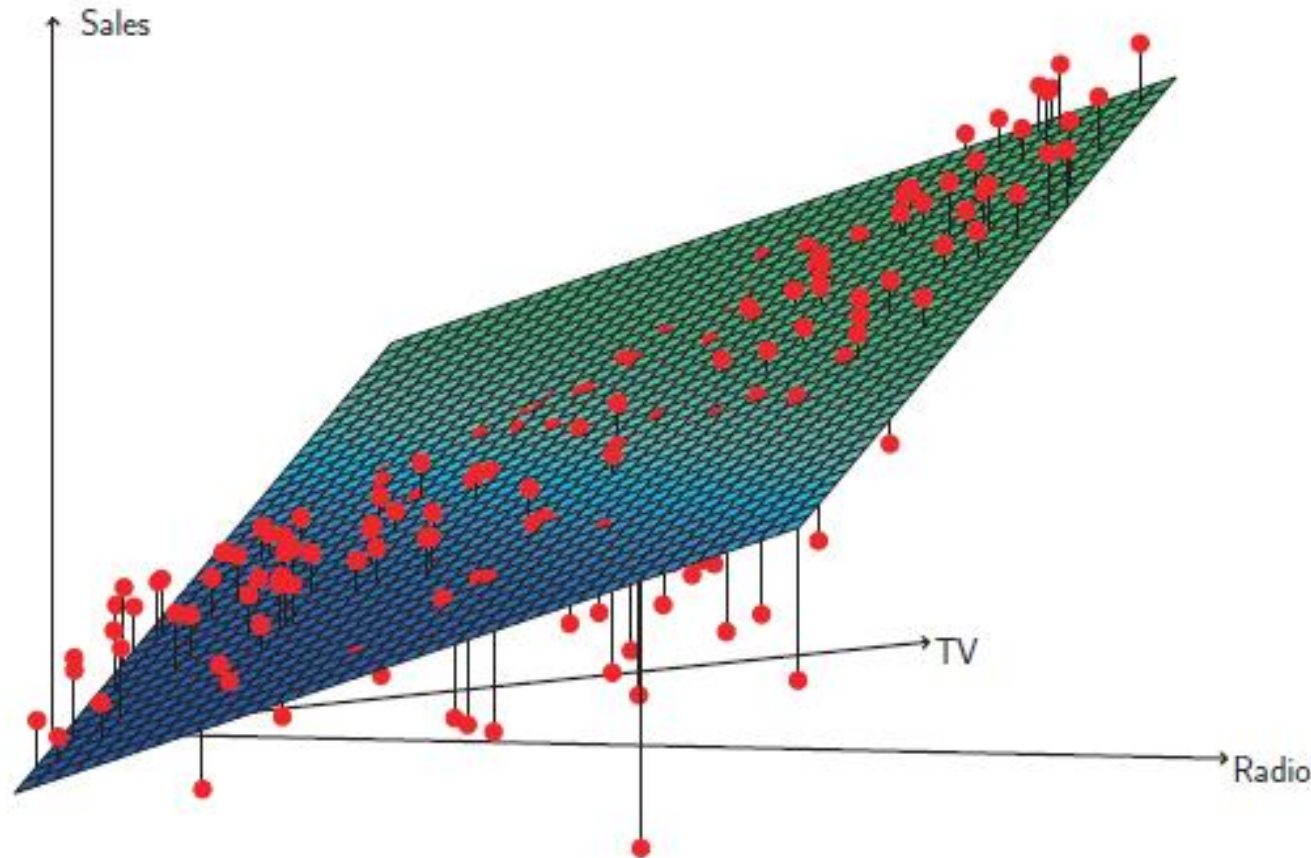


FIGURE 3.5. For the **Advertising** data, a linear regression fit to **sales** using **TV** and **radio** as predictors. From the pattern of the residuals, we can see that there is a pronounced non-linear relationship in the data. The positive residuals (those visible above the surface), tend to lie along the 45-degree line, where TV and Radio budgets are split evenly. The negative residuals (most not visible), tend to lie away from this line, where budgets are more lopsided.

- The least squares regression estimated from the data

$$\hat{Y} = \hat{\beta}_0 + \hat{\beta}_1 X_1 + \cdots + \hat{\beta}_p X_p$$

can differ from the population regression plane (model variance)

$$f(X) = \beta_0 + \beta_1 X_1 + \cdots + \beta_p X_p$$

- Moreover a linear model might not be enough to capture all the systematic components of the data (model bias)
- Finally the irreducible error is there ...

Example: Credit Predictors

- In the Credit dataset

- 7 quantitative
- 4 qualitative

- Quantitative ones

- Gender
- Student (status)
- Status (marital)
- Ethnicity
 - Caucasian
 - African American
 - Asian

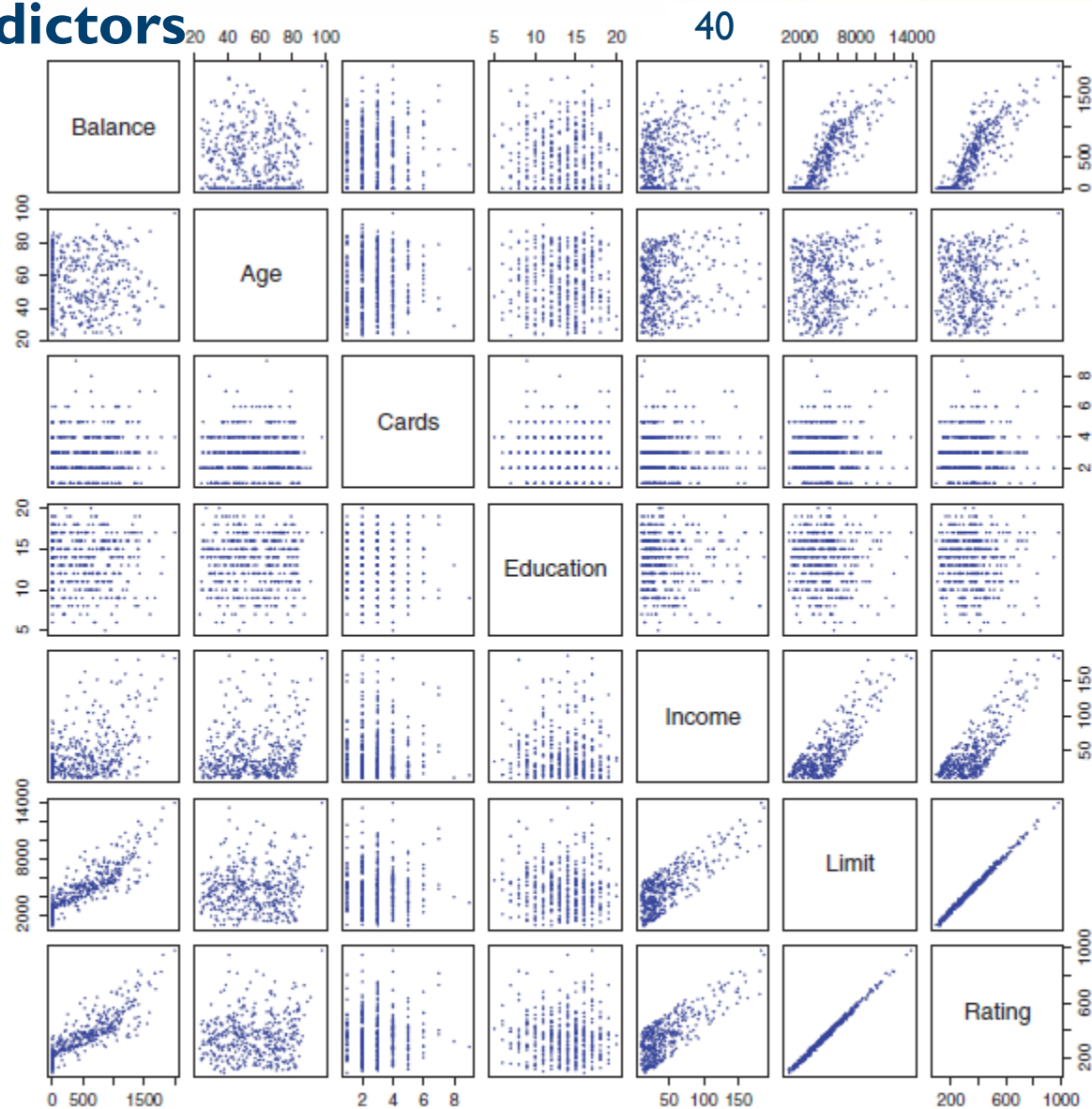


FIGURE 3.6. The **Credit** data set contains information about **balance**, **age**, **cards**, **education**, **income**, **limit**, and **rating** for a number of potential customers.

- Quantitative predictors with Two Levels can be coded using *Dummy Variables*

$$x_i = \begin{cases} 1 & \text{if } i\text{th person is female} \\ 0 & \text{if } i\text{th person is male,} \end{cases}$$

- This results in a “double” model for regression

$$y_i = \beta_0 + \beta_1 x_i + \epsilon_i = \begin{cases} \beta_0 + \beta_1 + \epsilon_i & \text{if } i\text{th person is female} \\ \beta_0 + \epsilon_i & \text{if } i\text{th person is male.} \end{cases}$$

Average balance
among males

Average difference of balance
between males and females

	Coefficient	Std. error	t-statistic	p-value
Intercept	509.80	33.13	15.389	< 0.0001
gender[Female]	19.73	46.05	0.429	0.6690

TABLE 3.7. *Least squares coefficient estimates associated with the regression of balance onto gender in the Credit data set. The linear model is given in (3.27). That is, gender is encoded as a dummy variable, as in (3.26).*

- Other possible coding can be devised for Dummy Variables

$$x_i = \begin{cases} 1 & \text{if } i\text{th person is female} \\ -1 & \text{if } i\text{th person is male} \end{cases}$$

- In this case the model becomes

$$y_i = \beta_0 + \beta_1 x_i + \epsilon_i = \begin{cases} \beta_0 + \beta_1 + \epsilon_i & \text{if } i\text{th person is female} \\ \beta_0 - \beta_1 + \epsilon_i & \text{if } i\text{th person is male.} \end{cases}$$

Average balance

Amount of salary females are above the average and males are below ...

- No significant impact on the regression output, but on the interpretation of the coefficients ...

- More than 2 levels are handled by using L-I dummy labels

$$x_{i1} = \begin{cases} 1 & \text{if } i\text{th person is Asian} \\ 0 & \text{if } i\text{th person is not Asian,} \end{cases}$$

$$x_{i2} = \begin{cases} 1 & \text{if } i\text{th person is Caucasian} \\ 0 & \text{if } i\text{th person is not Caucasian.} \end{cases}$$

- This again results in a “multiple output” model

$$y_i = \beta_0 + \beta_1 x_{i1} + \beta_2 x_{i2} + \epsilon_i = \begin{cases} \beta_0 + \beta_1 + \epsilon_i & \text{if } i\text{th person is Asian} \\ \beta_0 + \beta_2 + \epsilon_i & \text{if } i\text{th person is Caucasian} \\ \beta_0 + \epsilon_i & \text{if } i\text{th person is African American} \end{cases}$$

Average difference between African Americans and Asian

Average balance for African American

Average difference between African Americans and Caucasians

	Coefficient	Std. error	t-statistic	p-value
Intercept	531.00	46.32	11.464	< 0.0001
ethnicity[Asian]	−18.69	65.02	−0.287	0.7740
ethnicity[Caucasian]	−12.50	56.68	−0.221	0.8260

TABLE 3.8. Least squares coefficient estimates associated with the regression of **balance** onto **ethnicity** in the **Credit** data set. The linear model is given in (3.30). That is, **ethnicity** is encoded via two dummy variables (3.28) and (3.29).

- This again results in a “multiple output” model

The non coded level is defined baseline

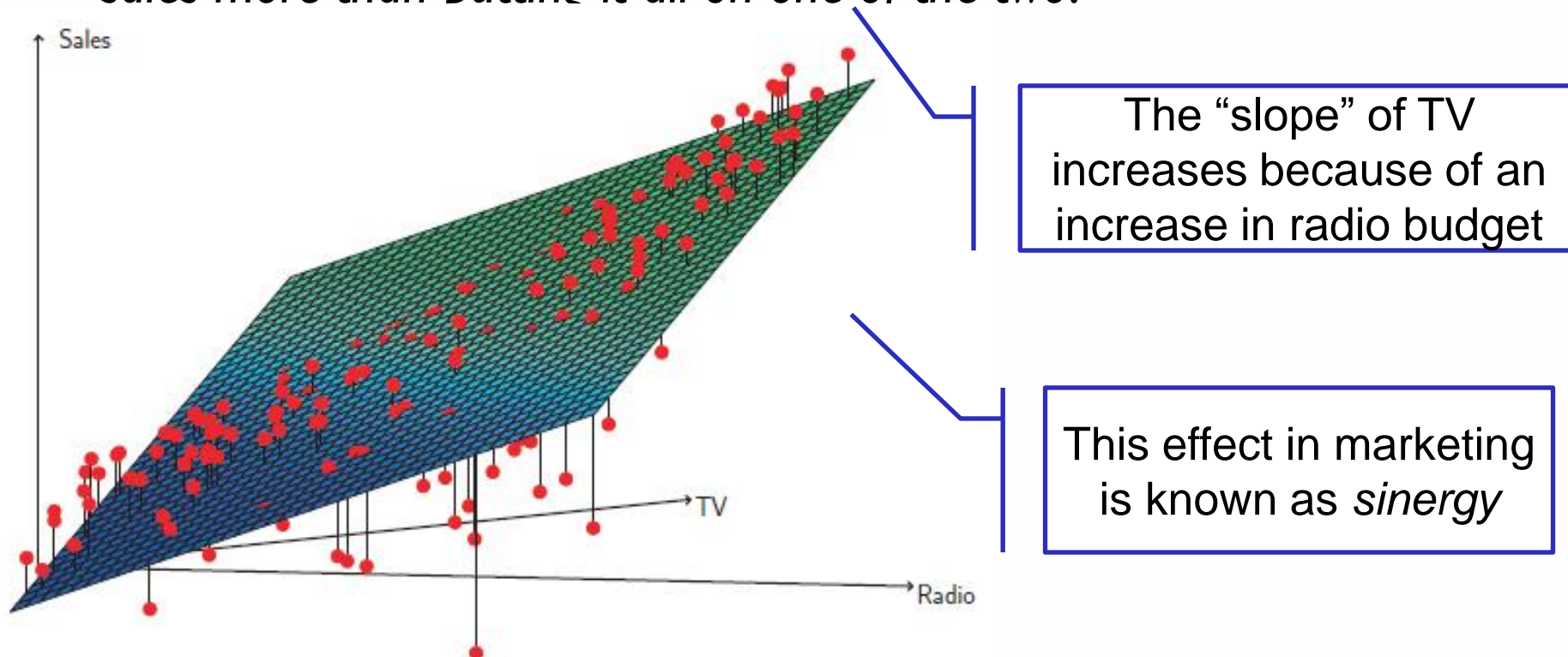
$$y_i = \beta_0 + \beta_1 x_{i1} + \beta_2 x_{i2} + \epsilon_i = \begin{cases} \beta_0 + \beta_1 + \epsilon_i & \text{if } i\text{th person is Asian} \\ \beta_0 + \beta_2 + \epsilon_i & \text{if } i\text{th person is Caucasian} \\ \beta_0 + \epsilon_i & \text{if } i\text{th person is African American} \end{cases}$$

- F-statistics should be used instead of p-values

$$H_0 : \beta_1 = \beta_2 = 0$$

F-statistics
p-value 0.96

- So far the linear regression model has assumed
 - Linear relationship between predictor and response
 - Additive relationship between predictor and response
- *What if allocating half the budget to TV and Radio would increase the sales more than putting it all on one of the two?*



- Let consider the classical Linear Regression model

$$Y = \beta_0 + \beta_1 X_1 + \beta_2 X_2 + \epsilon$$

- An increase in X_1 of 1 unit increases Y on average by β_1 units
- Presence or absence of other variables does not affect this

- We can extend the previous model with an interaction term

$$Y = \beta_0 + \beta_1 X_1 + \beta_2 X_2 + \beta_3 X_1 X_2 + \epsilon$$

- this translates in a “linear model”

One variable affects other variables influence

$$\begin{aligned} Y &= \beta_0 + (\beta_1 + \beta_3 X_2) X_1 + \beta_2 X_2 + \epsilon \\ &= \beta_0 + \tilde{\beta}_1 X_1 + \beta_2 X_2 + \epsilon \end{aligned}$$

- We can imagine some interaction in TV and Radio Advertising

$$\begin{aligned}\text{sales} &= \beta_0 + \beta_1 \times \text{TV} + \beta_2 \times \text{radio} + \beta_3 \times (\text{radio} \times \text{TV}) + \epsilon \\ &= \beta_0 + (\beta_1 + \beta_3 \times \text{radio}) \times \text{TV} + \beta_2 \times \text{radio} + \epsilon.\end{aligned}$$

- The interaction term is the increase of effectiveness of TV advertising for one unit of Radio advertising
- p-value suggests this interaction to be significant
- R^2 increases from 89.7% to 96.8% (69% of missing variance)

	Coefficient	Std. error	t-statistic	p-value
Intercept	6.7502	0.248	27.23	< 0.0001
TV	0.0191	0.002	12.70	< 0.0001
radio	0.0289	0.009	3.24	0.0014
TV×radio	0.0011	0.000	20.73	< 0.0001

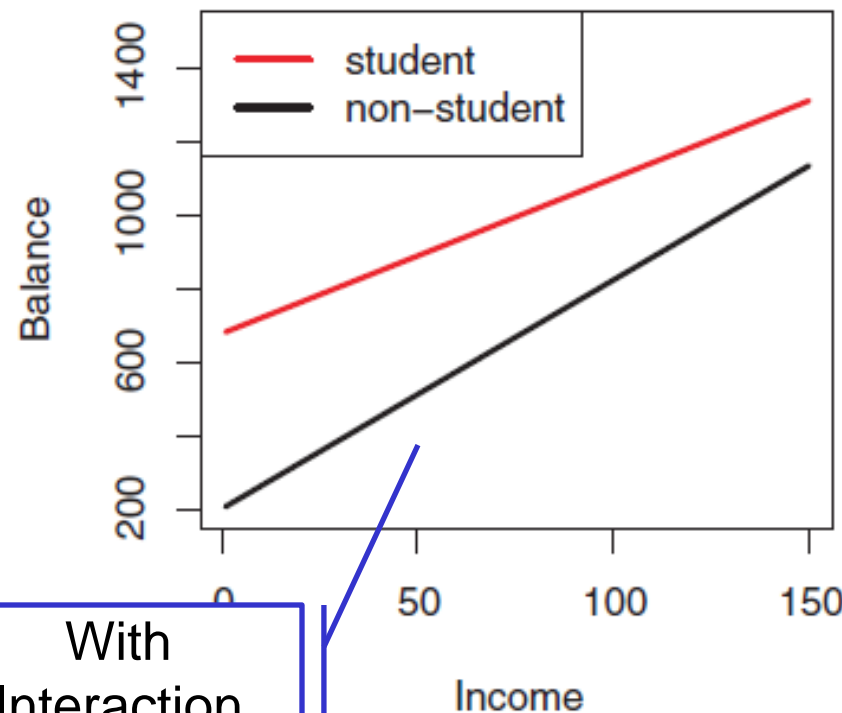
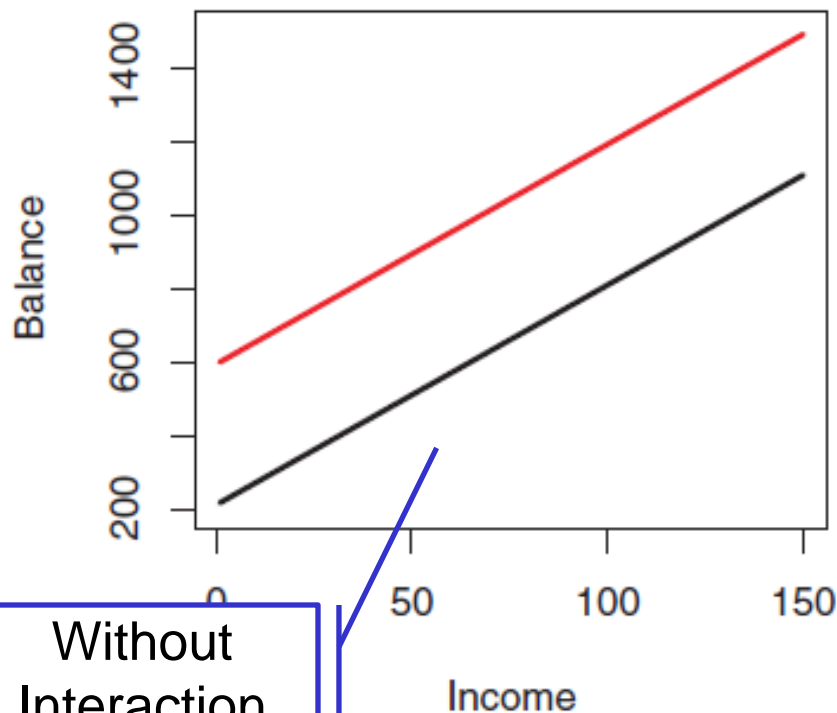
TABLE 3.9. For the Advertising data, least squares coefficient estimates associated with the regression of sales onto TV and radio, with an interaction term, as in (3.33).

- We can check for interactions with qualitative predictors

$$\text{balance}_i \approx \beta_0 + \beta_1 \times \text{income}_i + \begin{cases} \beta_2 + \beta_3 \times \text{income}_i & \text{if student} \\ \beta_2 & \text{if not student} \end{cases}$$

$$= \begin{cases} (\beta_0 + \beta_2 + \beta_3 \times \text{income}_i) & \text{if student} \\ \beta_0 + \beta_1 \times \text{income}_i & \text{if not student} \end{cases}$$

Is this term significant?



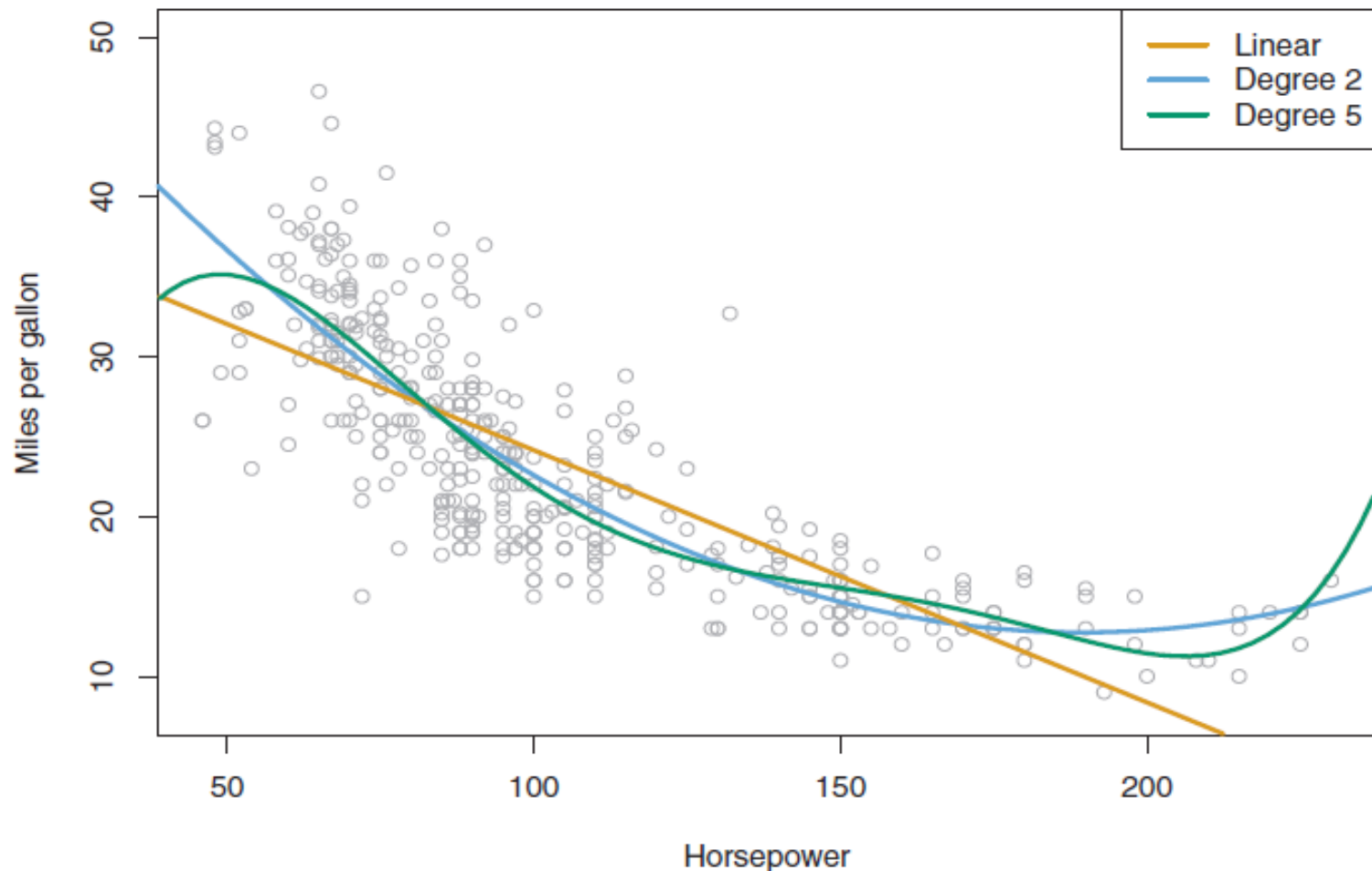


FIGURE 3.8. The **Auto** data set. For a number of cars, **mpg** and **horsepower** are shown. The linear regression fit is shown in orange. The linear regression fit for a model that includes **horsepower**² is shown as a blue curve. The linear regression fit for a model that includes all polynomials of **horsepower** up to fifth-degree is shown in green.

- We can use Polynomial Regression to accomodate non linearity

$$\text{mpg} = \beta_0 + \beta_1 \times \text{horsepower} + \beta_2 \times \text{horsepower}^2 + \epsilon$$

- It is still a linear fitting problem !!!!
- A 5th grade polynome seems too much, but the quadratic term is statistically significant

	Coefficient	Std. error	t-statistic	p-value
Intercept	56.9001	1.8004	31.6	< 0.0001
horsepower	−0.4662	0.0311	−15.0	< 0.0001
horsepower ²	0.0012	0.0001	10.1	< 0.0001

TABLE 3.10. For the **Auto** data set, least squares coefficient estimates associated with the regression of **mpg** onto **horsepower** and **horsepower²**.

- A number of possible problems might be encountered when fitting the linear regression model.
 - Non-linearity of the data
 - Dependence of the error terms
 - Non-constant variance of error terms
 - Outliers
 - High leverage points
 - Collinearity

In practice, identifying and overcoming these problems is as much an art as a science. Many pages in countless books have been written on this topic. Since the linear regression model is not our primary focus here, we will provide only a brief summary of some key points.

- If the linearity assumption does not hold, the conclusions drawn might be (at least) inaccurate ... check residual plot!

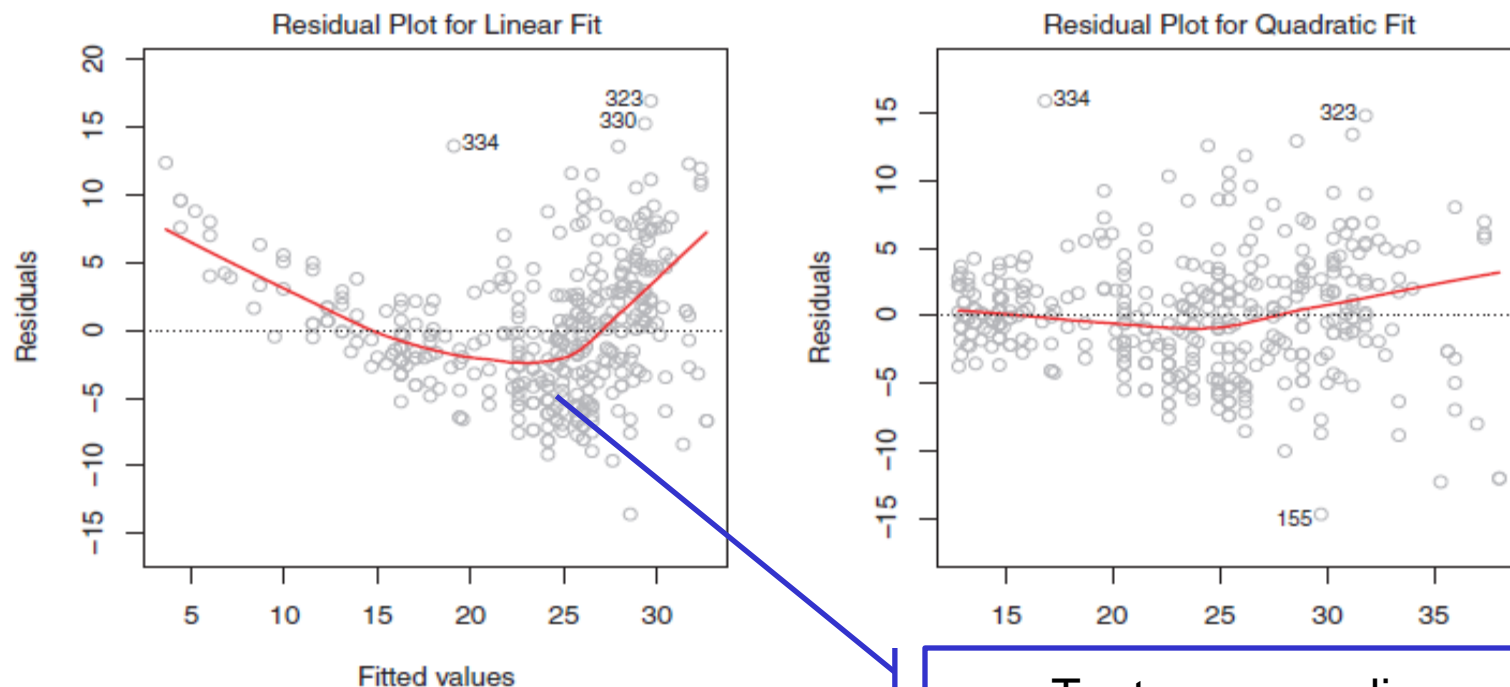


FIGURE 3.9. Plots of residuals versus predicted values for a data set. In each plot, the red line is a smooth fit to the residuals to make it easier to identify a trend. Left: A linear regression of `mpg` on `horsepower`. A strong pattern in the residuals indicates non-linearity in the data. Right: A linear regression of `mpg` on `horsepower` and `horsepower2`. There is little pattern in the residuals.

Try to use non linear transformations of the predictor (log X, X², sqrt(X), ...)

- Errors are supposed to uncorrelated otherwise standard errors would underestimate the true errors ... tracking phenomenon

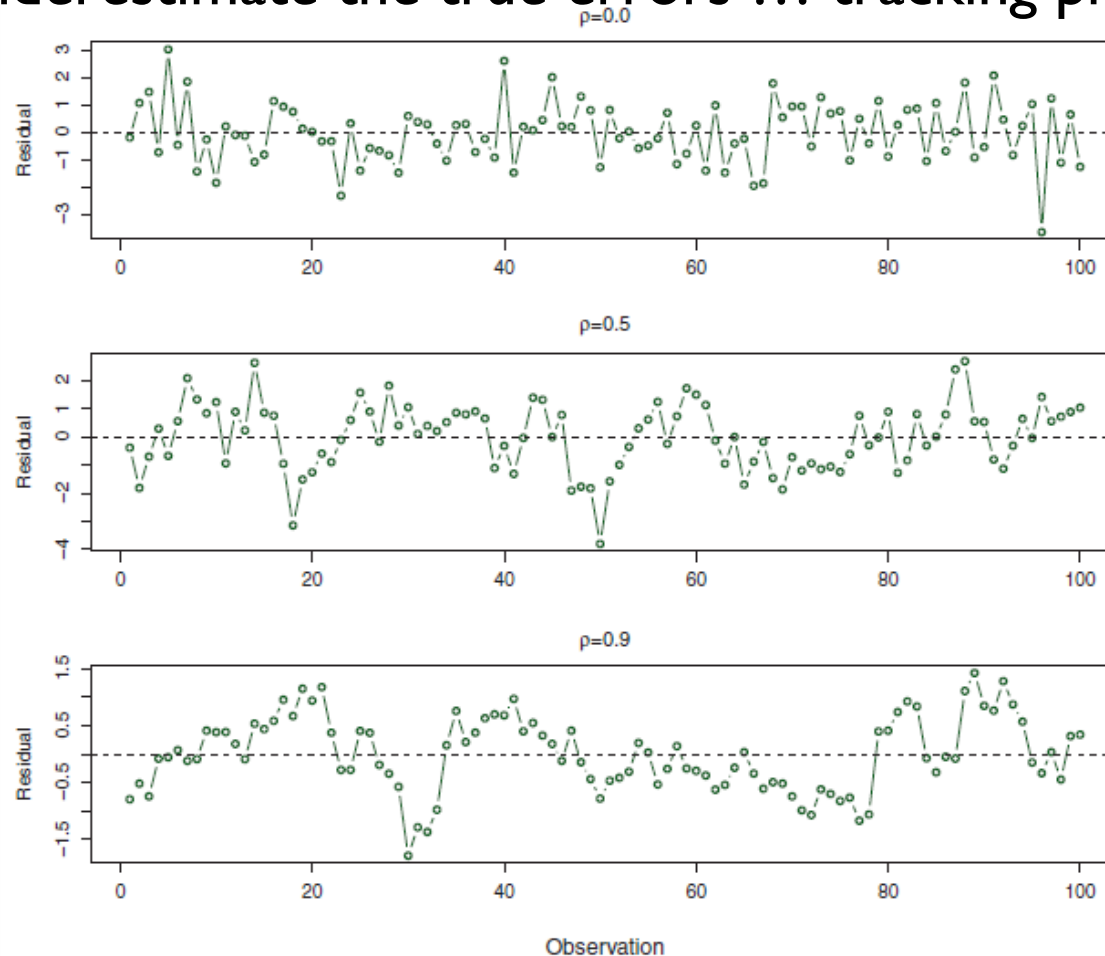


FIGURE 3.10. Plots of residuals from simulated time series data sets generated with differing levels of correlation ρ between error terms for adjacent time points.

- Linear Regression assumes no heteroscedasticity in the noise

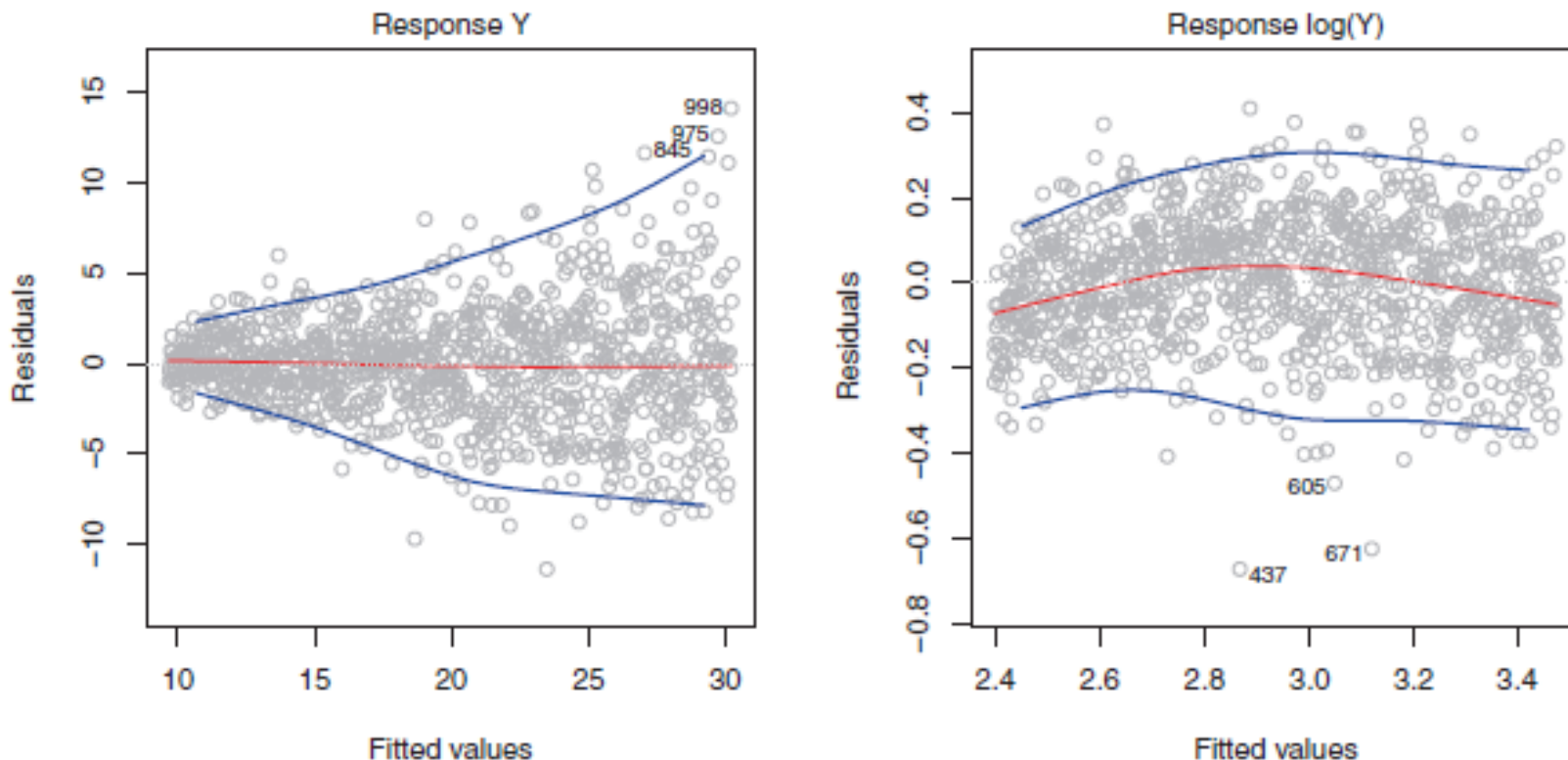


FIGURE 3.11. *Residual plots. In each plot, the red line is a smooth fit to the residuals, intended to make it easier to identify a trend. The blue lines track the outer quantiles of the residuals, and emphasize patterns. Left: The funnel shape indicates heteroscedasticity. Right: The predictor has been log-transformed, and there is now no evidence of heteroscedasticity.*

- An outlier is a point that is too far from its prediction
 - Might be due to error in data collection (just remove it)
 - Might be due to some missing predictors (revise the model)

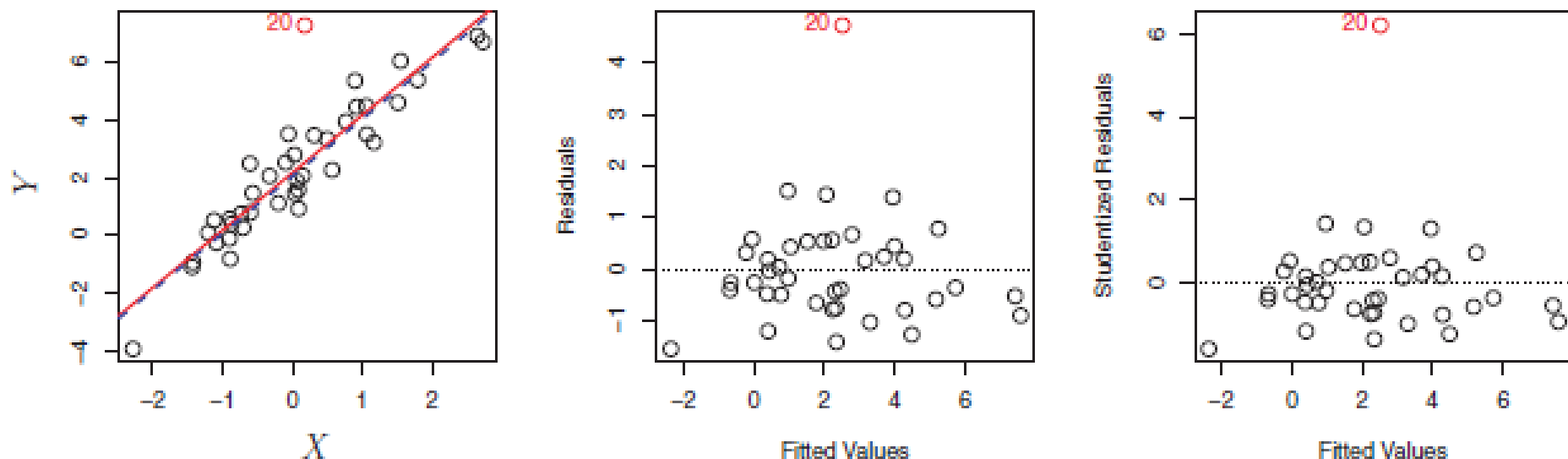


FIGURE 3.12. Left: The least squares regression line is shown in red, and the regression line after removing the outlier is shown in blue. Center: The residual plot clearly identifies the outlier. Right: The outlier has a studentized residual of 6; typically we expect values between -3 and 3 .

- High leverage points have unexpected values for a predictor

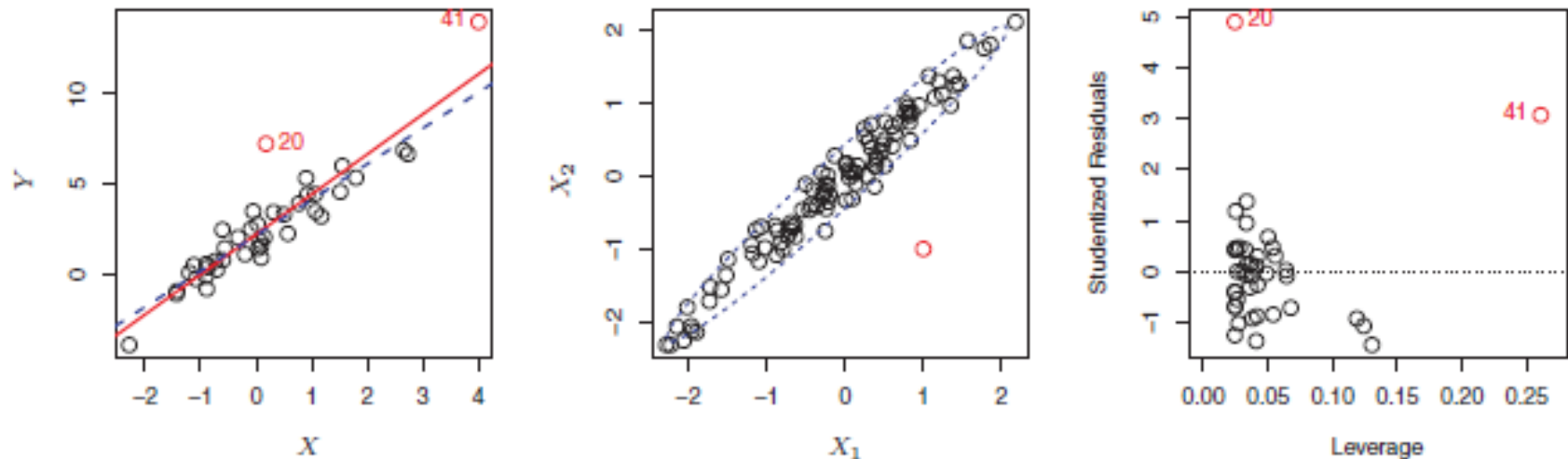


FIGURE 3.13. Left: Observation 41 is a high leverage point, while 20 is not. The red line is the fit to all the data, and the blue line is the fit with observation 41 removed. Center: The red observation is not unusual in terms of its X_1 value or its X_2 value, but still falls outside the bulk of the data, and hence has high leverage. Right: Observation 41 has a high leverage and a high residual.

- Leverage statistics (between $1/n$ and 1, average $(p+1)/n$)

$$h_i = \frac{1}{n} + \frac{(x_i - \bar{x})^2}{\sum_{i'=1}^n (x_{i'} - \bar{x})^2}$$

- Some factors might be highly related so it might be difficult to separate their effects

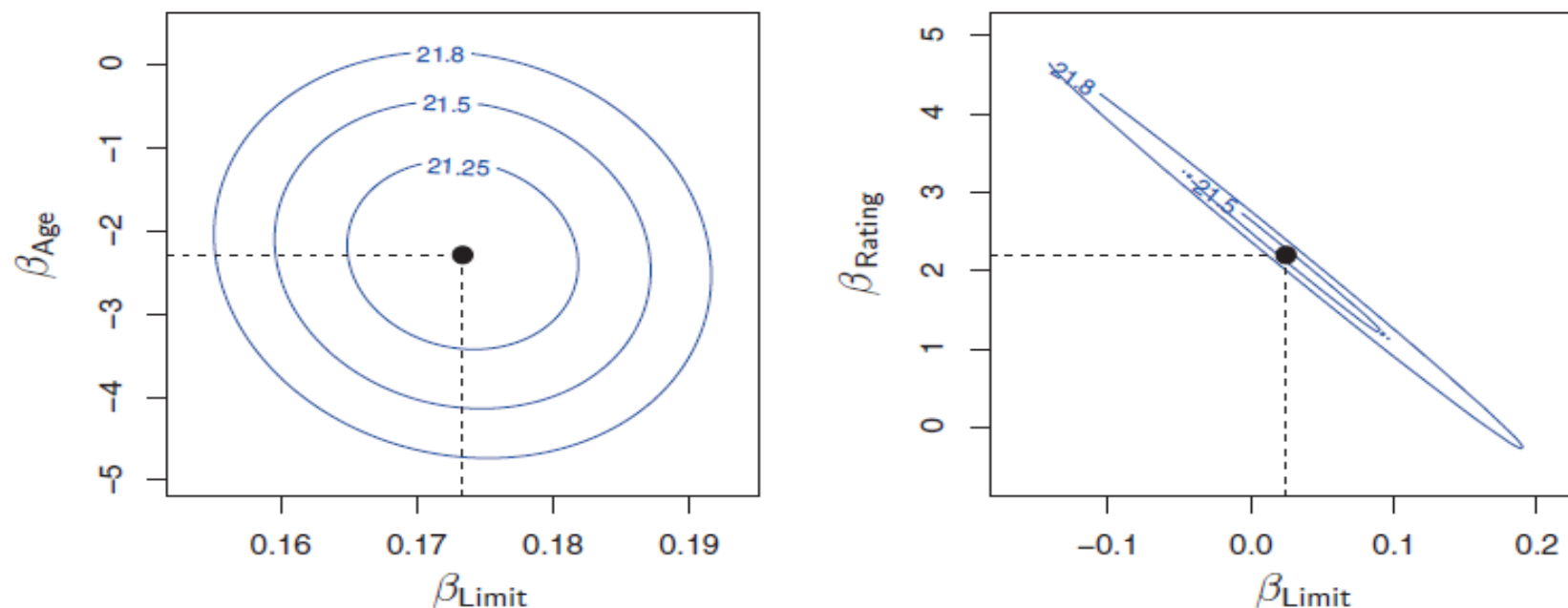
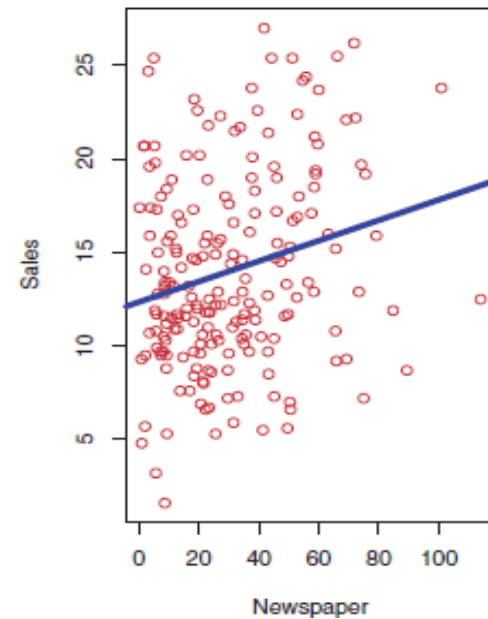
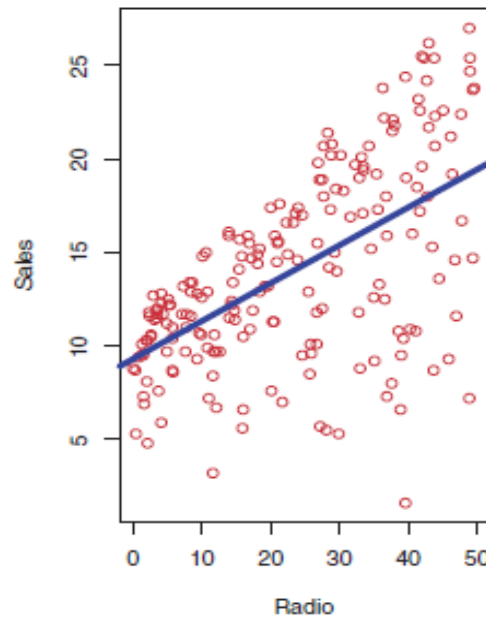
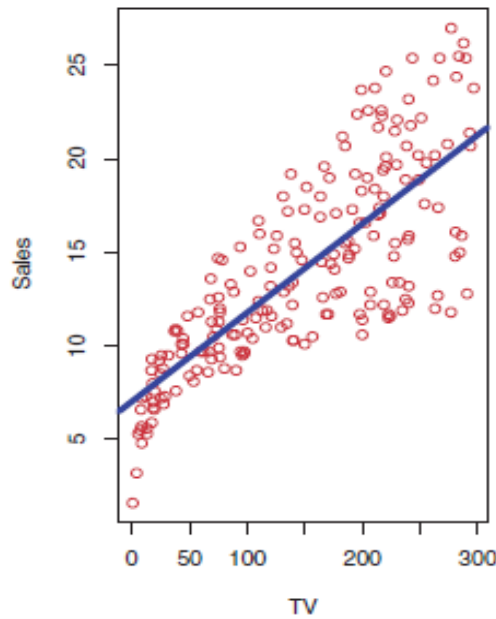


FIGURE 3.15. Contour plots for the RSS values as a function of the parameters β for various regressions involving the **Credit** data set. In each plot, the black dots represent the coefficient values corresponding to the minimum RSS. Left: A contour plot of RSS for the regression of **balance** onto **age** and **limit**. The minimum value is well defined. Right: A contour plot of RSS for the regression of **balance** onto **rating** and **limit**. Because of the collinearity, there are many pairs $(\beta_{\text{Limit}}, \beta_{\text{Rating}})$ with a similar value for RSS.



- *Is there a relationship between advertising budget and sales?*
- *How strong is the relationship between advertising budget and sales?*
- *Which media contribute to sales?*
- *How accurately can we estimate the effect of each medium on sales?*
- *How accurately can we predict future sales?*
- *Is the relationship linear?*
- *Is there synergy among the advertising media?*

- Simple Linear Regression Model
- Multi Variate Regression Model
- Some Considerations in Regression Model
- Feature subset selection
 - Best subset selection
 - Stepwise selection
 - Choosing the optimal model
- Shrinkage methods
 - Ridge Regression
 - The Lasso

- We can devise alternative procedures to least squares
 - Improve prediction accuracy: if number of data is limited (or p is big) we might have “low bias” but too “high variance” (overfitting) and a poor prediction
 - Improve model interpretability: irrelevant variables, beside impacting on prediction accuracy, make the model unnecessary complex and difficult to interpret
- Several alternatives to remove unnecessary features (predictors)
 - Subset Selection: selection of the input variables
 - Shrinkage (or regularization): reduction of model variance
 - Dimension reduction: projection on an input subspace

- Select the variables which are really associated to the prediction
 - Exhaustive exploration of model space (2^p)
 - Forward selection
 - Backward selection
 - Mixed selection
- Exhaustive exploration is unfeasible because of exponential complexity
 - $Y =$
 - $Y =$
 - $Y =$
 - $Y =$

If $p=30$ the number of possible models is 1.073.741.824

Different possible metrics, e.g., C_p , AIC, BIC, adjusted R^2

- Fit a least squares regression for any possible input combination
 - A total of 2^p need to be compared
 - Best Subset Selection introduces a procedure to evaluate them systematically

Algorithm 6.1 *Best subset selection*

1. Let \mathcal{M}_0 denote the *null model*, which contains no predictors. This model simply predicts the sample mean for each observation.
2. For $k = 1, 2, \dots, p$:
 - (a) Fit all $\binom{p}{k}$ models that contain exactly k predictors.
 - (b) Pick the best among these $\binom{p}{k}$ models, and call it \mathcal{M}_k . Here *best* is defined as having the smallest RSS, or equivalently largest R^2 .
3. Select a single best model from validated prediction error, C_p (AIC), BIC , or adjusted R^2 .

Cross-validated prediction error!

 using cross-validated prediction error, C_p (AIC), BIC , or adjusted R^2 .

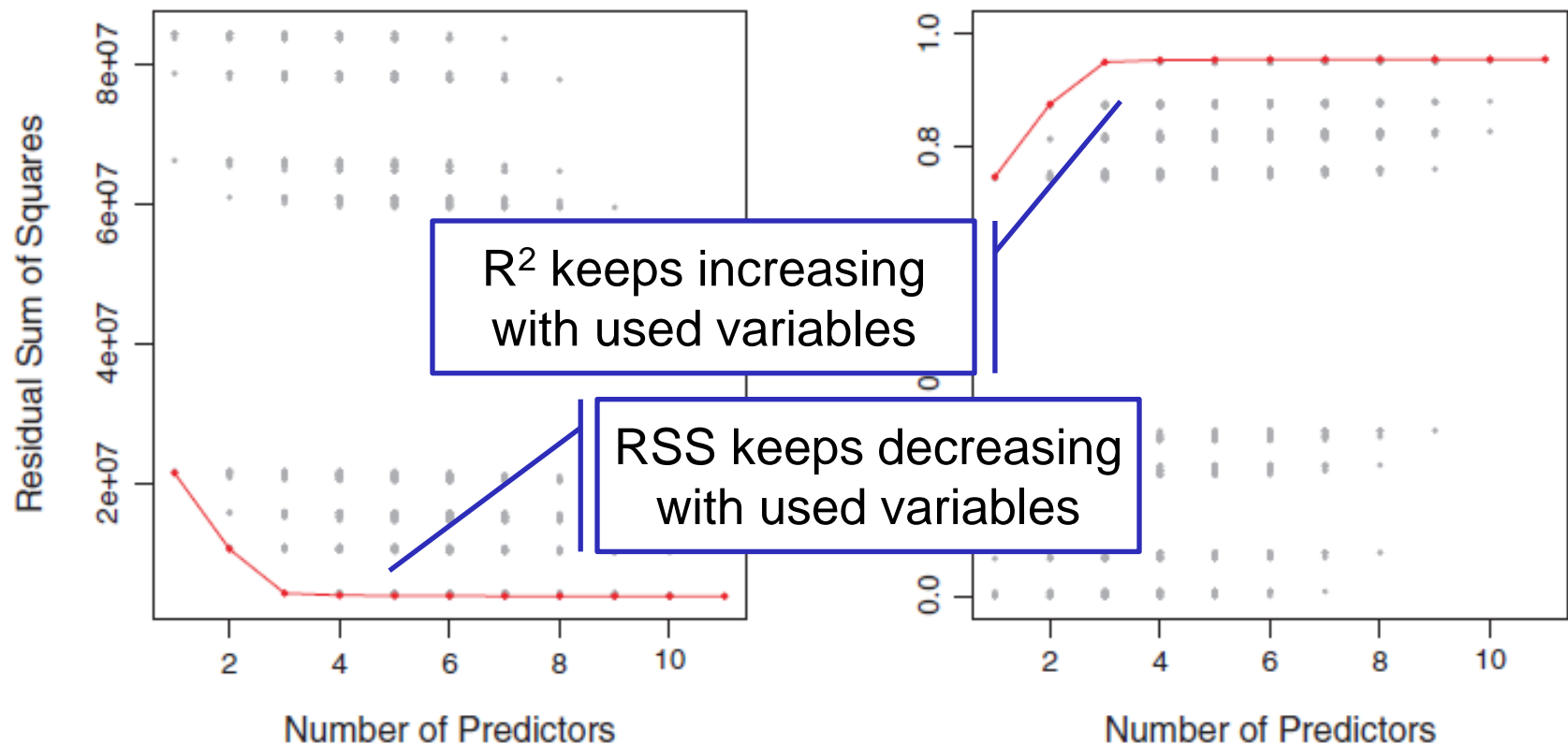


FIGURE 6.1. For each possible model containing a subset of the ten predictors in the **Credit** data set, the RSS and R^2 are displayed. The red frontier tracks the best model for a given number of predictors, according to RSS and R^2 . Though the data set contains only ten predictors, the x -axis ranges from 1 to 11, since one of the variables is categorical and takes on three values, leading to the creation of two dummy variables.

- *Forward stepwise selection* is a computationally efficient alternative
 - Starts from an empty model with no predictors
 - Adds one predictor at the time until “all are in”
 - At each stage adds “most improving” variable

Algorithm 6.2 *Forward stepwise selection*

1. Let \mathcal{M}_0 denote the *null* model, which contains no predictors.
 2. For $k = 0, \dots, p - 1$:
 - (a) Consider all $p - k$ models that augment the predictors in \mathcal{M}_k with one additional predictor.
 - (b) Choose the *best* among these $p - k$ models, and call it \mathcal{M}_{k+1} . Here *best* is defined as having smallest RSS or highest R^2 .
 3. Select a single best model from among $\mathcal{M}_0, \dots, \mathcal{M}_p$ using cross-validated prediction error, C_p (AIC), BIC, or adjusted R^2 .
-

- Forward Stepwise is a greedy approach
 - Needs to fit $1 + p(p+1)/2$ models instead of 2^p
 - It can be used also when $n < p$ (it will stop with $k < n$ variables)
 - It does not “reconsider” its choices and thus might result in a suboptimal subset of variables

# Variables	Best subset	Forward stepwise
One	rating	rating
Two	rating, income	rating, income
Three	rating, income, student	rating, income, student
Four	cards, income student, limit	rating, income, student, limit

TABLE 6.1. *The first four selected models for best subset selection and forward stepwise selection on the **Credit** data set. The first three models are identical but the fourth models differ.*

- *Backward stepwise selection* is yet computationally efficient
 - Starts from the model having all predictors
 - At each stage removes the “least useful” variable

Algorithm 6.3 *Backward stepwise selection*

1. Let \mathcal{M}_p denote the *full* model, which contains all p predictors.
2. For $k = p, p - 1, \dots, 1$:
 - (a) Consider all k models that contain all but one of the predictors in \mathcal{M}_k , for a total of $k - 1$ predictors.
 - (b) Choose the *best* among these k models, and call it \mathcal{M}_{k-1} . Here *best* is defined as having smallest RSS or highest R^2 .
3. Select a single best model from among $\mathcal{M}_0, \dots, \mathcal{M}_p$ using cross-validated prediction error, C_p (AIC), BIC, or adjusted R^2 .

-
- Greedy as Forward Stepwise, but cannot be used when $n < p$

- Feature subset selection algorithms “optimize” the number of features according to RSS and R^2 , but on the test set
- Several approaches estimate test error correct the training error

Strong statistical background

- Mallows CP

$$C_p = \frac{1}{n} (\text{RSS} + 2d\hat{\sigma}^2)$$

d = number of predictors

- Akaike Information Criterion

$$\text{AIC} = \frac{1}{n\hat{\sigma}^2} (\text{RSS} + 2d\hat{\sigma}^2)$$

$\hat{\sigma}^2$ = estimate of the variance associated to the complete model

- Bayesian Information Criterion

$$\text{BIC} = \frac{1}{n} (\text{RSS} + \log(n)d\hat{\sigma}^2)$$

Some constants omitted, but proportional to C_p

- Adjusted R^2

$$\text{Adjusted } R^2 = 1 - \frac{\text{RSS}/(n - d - 1)}{\text{TSS}/(n - 1)}$$

Some constants omitted, more stringent than C_p

Equivalent to $\frac{\text{RSS}}{n-d-1}$

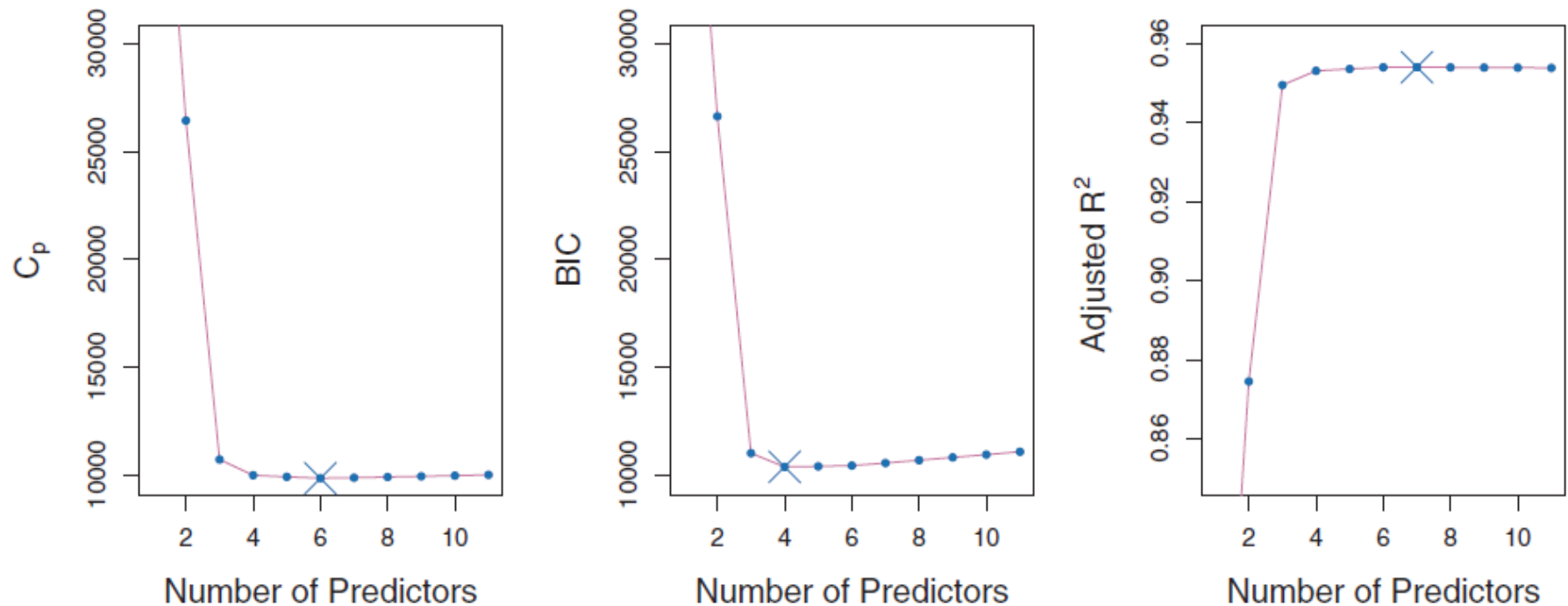


FIGURE 6.2. C_p , BIC , and adjusted R^2 are shown for the best models of each size for the **Credit** data set (the lower frontier in Figure 6.1). C_p and BIC are estimates of test MSE. In the middle plot we see that the BIC estimate of test error shows an increase after four variables are selected. The other two plots are rather flat after four variables are included.

- We can use data itself to estimate the error on new data
 - We can use an hold out set and perform validation

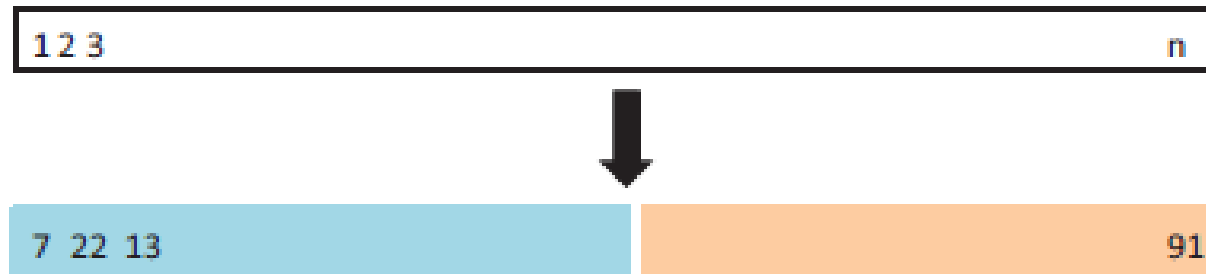


FIGURE 5.1. *A schematic display of the validation set approach. A set of n observations are randomly split into a training set (shown in blue, containing observations 7, 22, and 13, among others) and a validation set (shown in beige, and containing observation 91, among others). The statistical learning method is fit on the training set, and its performance is evaluated on the validation set.*

- We can use data itself to estimate the error on new data
 - We can use an hold out set and perform validation
 - We can use k-fold cross-validation

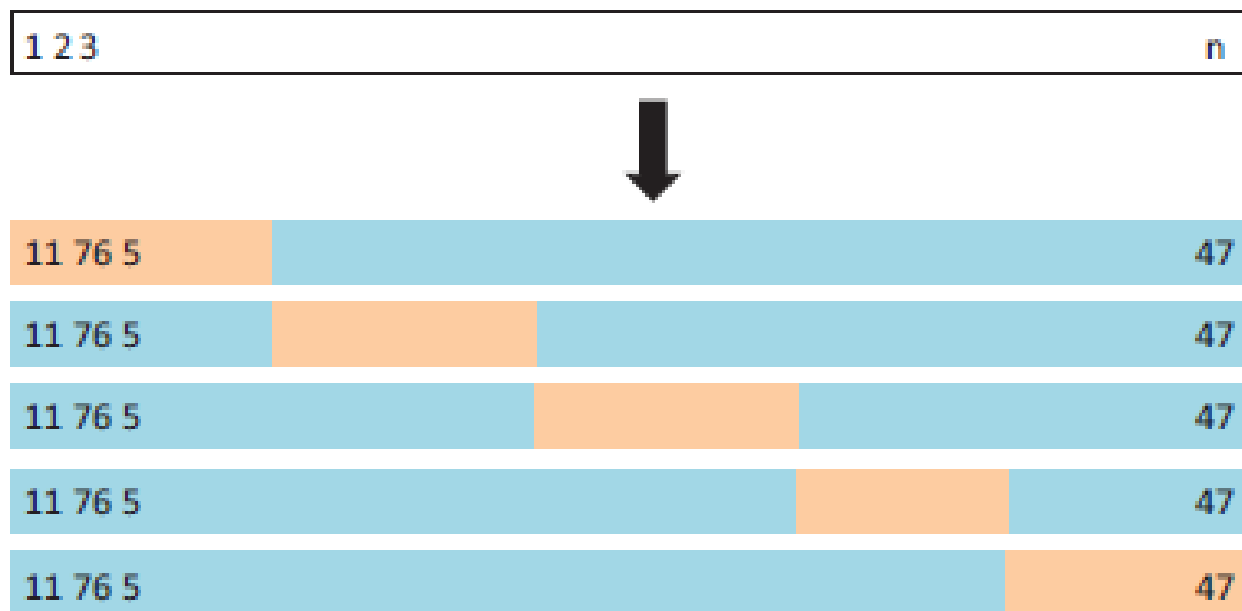


FIGURE 5.5. A schematic display of 5-fold CV. A set of n observations is randomly split into five non-overlapping groups. Each of these fifths acts as a validation set (shown in beige), and the remainder as a training set (shown in blue). The test error is estimated by averaging the five resulting MSE estimates.

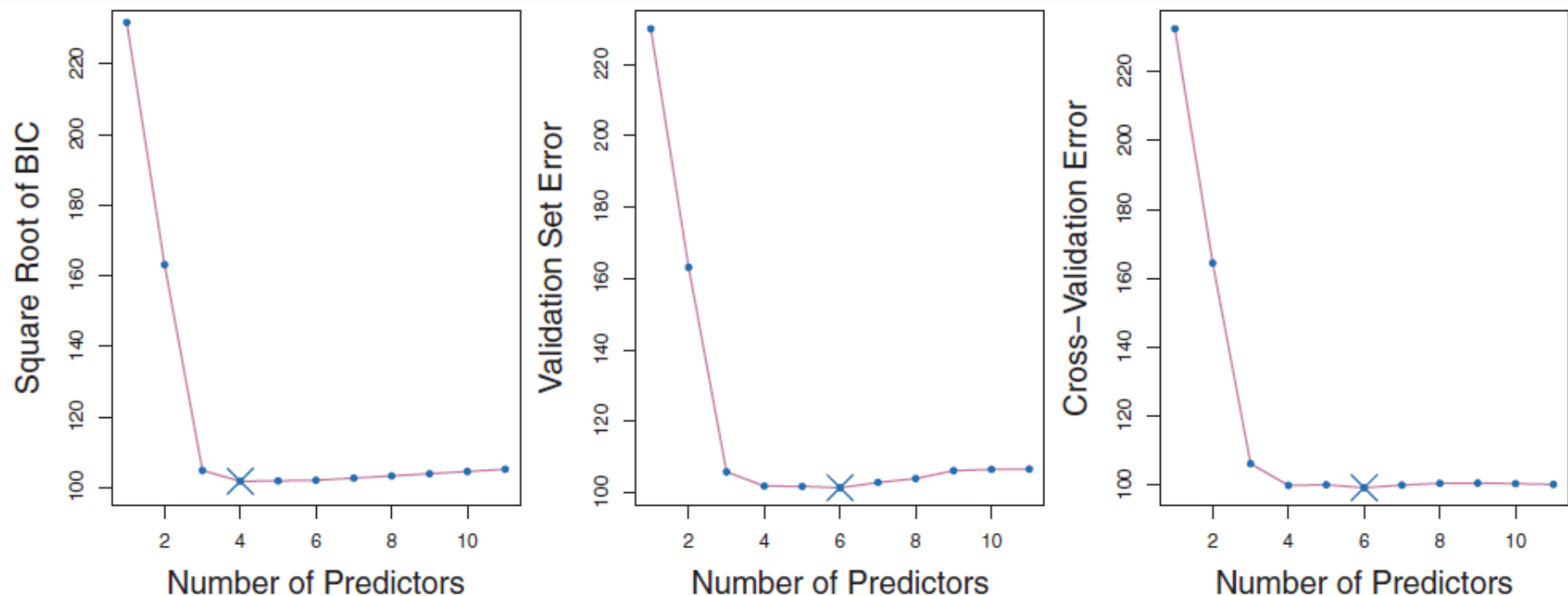


FIGURE 6.3. For the **Credit** data set, three quantities are displayed for the best model containing d predictors, for d ranging from 1 to 11. The overall best model, based on each of these quantities, is shown as a blue cross. Left: Square root of BIC. Center: Validation set errors. Right: Cross-validation errors.

- Ordinary Least Squares (OLS) minimizes

$$\text{RSS} = \sum_{i=1}^n \left(y_i - \beta_0 - \sum_{j=1}^p \beta_j x_{ij} \right)^2$$

- Ridge Regression minimizes a slightly different function

$$\sum_{i=1}^n \left(y_i - \beta_0 - \sum_{j=1}^p \beta_j x_{ij} \right)^2 + \lambda \sum_{j=1}^p \beta_j^2 = \text{RSS} + \lambda \sum_{j=1}^p \beta_j^2$$

- $\lambda \geq 0$ is a tuning parameter to be estimated experimentally
- Shrinkage does not apply to intercept, with centered variables

$$\hat{\beta}_0 = \bar{y} = \sum_{i=1}^n y_i / n$$

- $\lambda \sum_j \beta_j^2$ is called shrinkage penalty
- as $\lambda \rightarrow \infty$ parameters shrink to zero

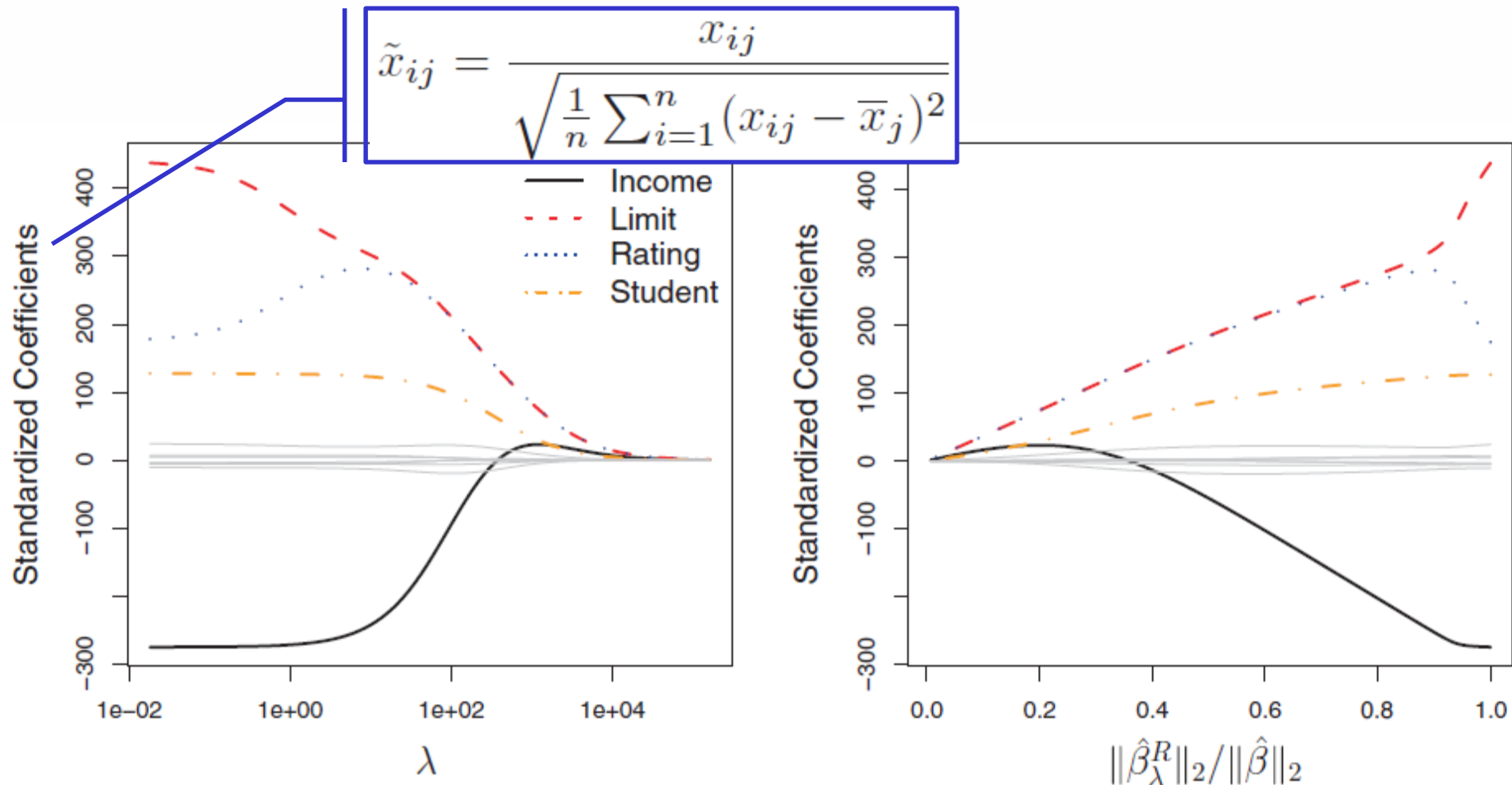


FIGURE 6.4. The standardized ridge regression coefficients are displayed for the **Credit** data set, as a function of λ and $\|\hat{\beta}_\lambda^R\|_2 / \|\hat{\beta}\|_2$.

$$\|\beta\|_2 = \sqrt{\sum_{j=1}^p \beta_j^2}$$

- Ridge regression improves over OLS because of a reduced model variance (i.e., a better bias-variance trade-off)

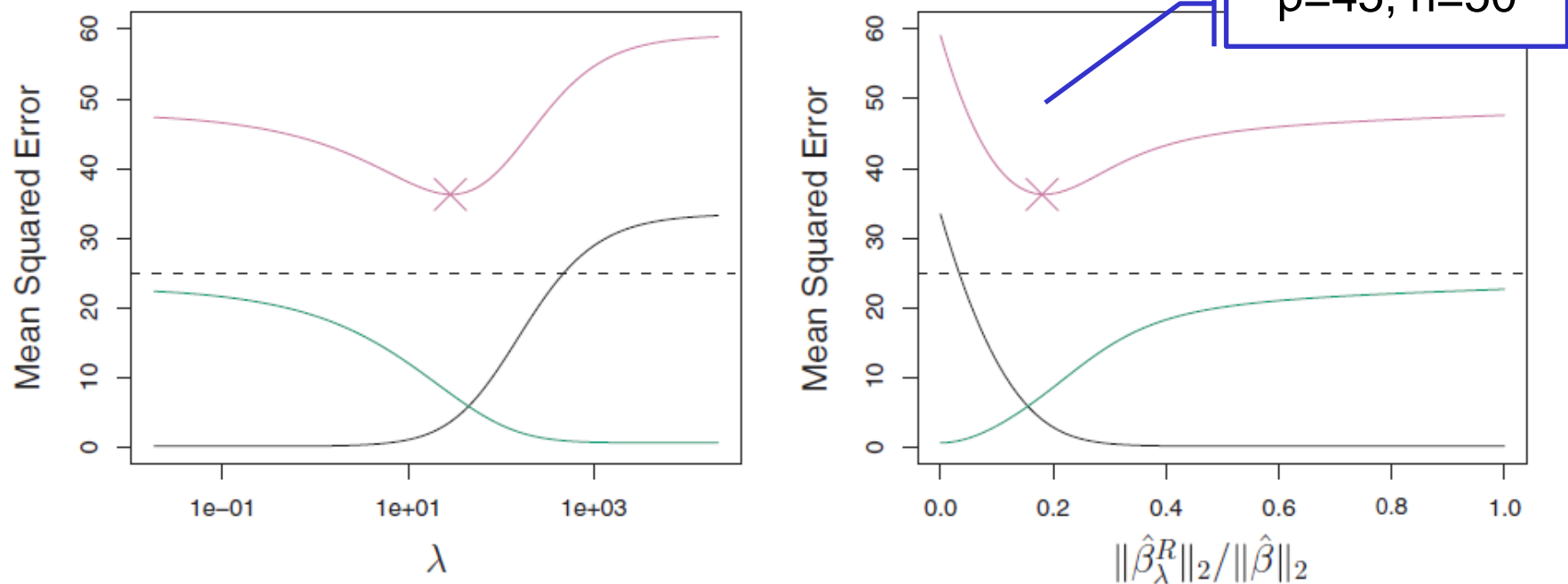


FIGURE 6.5. Squared bias (black), variance (green), and test mean squared error (purple) for the ridge regression predictions on a simulated data set, as a function of λ and $\|\hat{\beta}_\lambda^R\|_2 / \|\hat{\beta}\|_2$. The horizontal dashed lines indicate the minimum possible MSE. The purple crosses indicate the ridge regression models for which the MSE is smallest.

- Ridge regression is much more efficient than subset selection, however it will use all the p input
- The Lasso is an alternative to shrink regression coefficients

$$\sum_{i=1}^n \left(y_i - \beta_0 - \sum_{j=1}^p \beta_j x_{ij} \right)^2 + \lambda \sum_{j=1}^p |\beta_j| = \text{RSS} + \lambda \sum_{j=1}^p |\beta_j|$$

- The $\|\beta\|_1 = \sum |\beta_j|$ forces coefficients to be exactly zero
- The Lasso performs *variable selection*
- Models are simpler, sparse, and easy to interpret

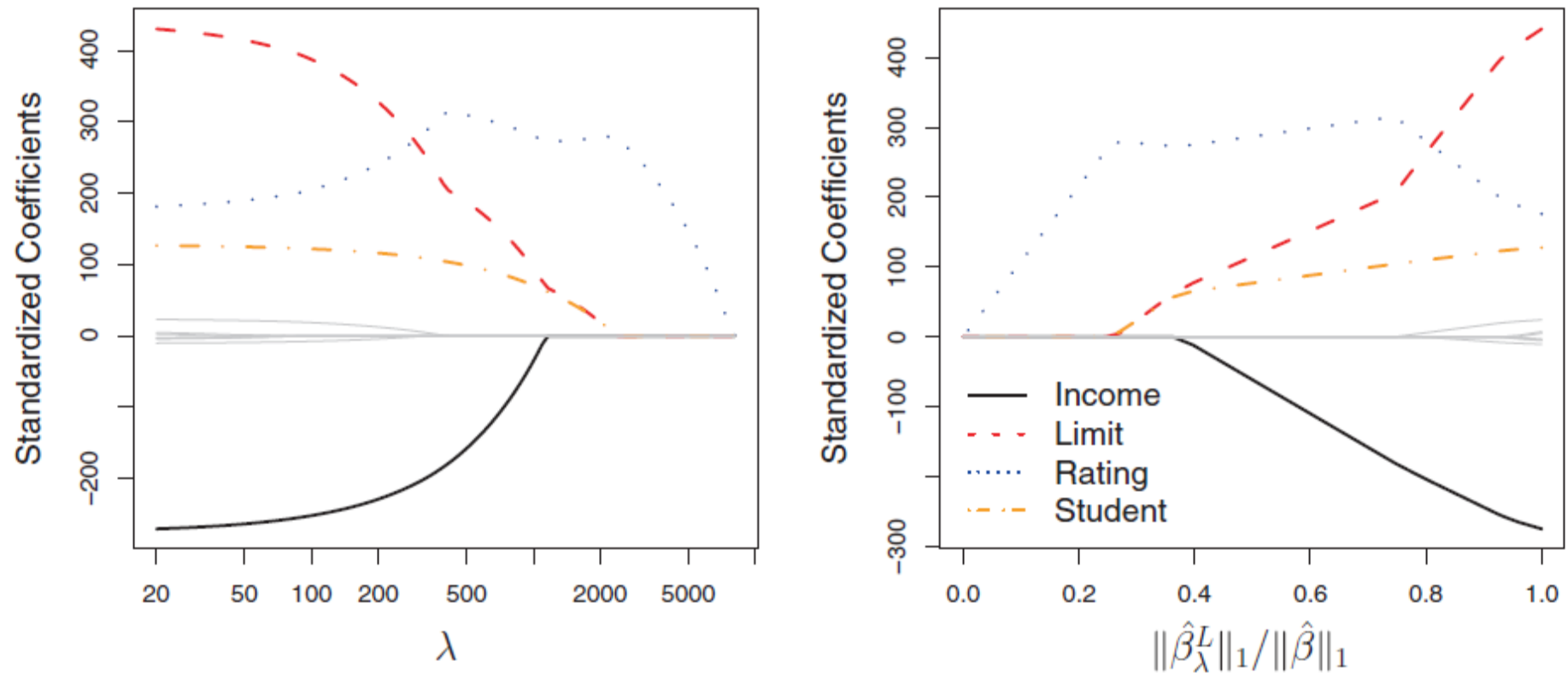


FIGURE 6.6. The standardized lasso coefficients on the **Credit** data set are shown as a function of λ and $\|\hat{\beta}_\lambda^L\|_1 / \|\hat{\beta}\|_1$.

- We can show that Ridge regression solves the problem

$$\underset{\beta}{\text{minimize}} \left\{ \sum_{i=1}^n \left(y_i - \beta_0 - \sum_{j=1}^p \beta_j x_{ij} \right)^2 \right\} \quad \text{subject to} \quad \sum_{j=1}^p \beta_j^2 \leq s$$

- While The Lasso solves the problem

$$\underset{\beta}{\text{minimize}} \left\{ \sum_{i=1}^n \left(y_i - \beta_0 - \sum_{j=1}^p \beta_j x_{ij} \right)^2 \right\} \quad \text{subject to} \quad \sum_{j=1}^p |\beta_j| \leq s$$

- They approximate the Best Subset Selection

$$\underset{\beta}{\text{minimize}} \left\{ \sum_{i=1}^n \left(y_i - \beta_0 - \sum_{j=1}^p \beta_j x_{ij} \right)^2 \right\} \quad \text{subject to} \quad \sum_{j=1}^p I(\beta_j \neq 0) \leq s$$

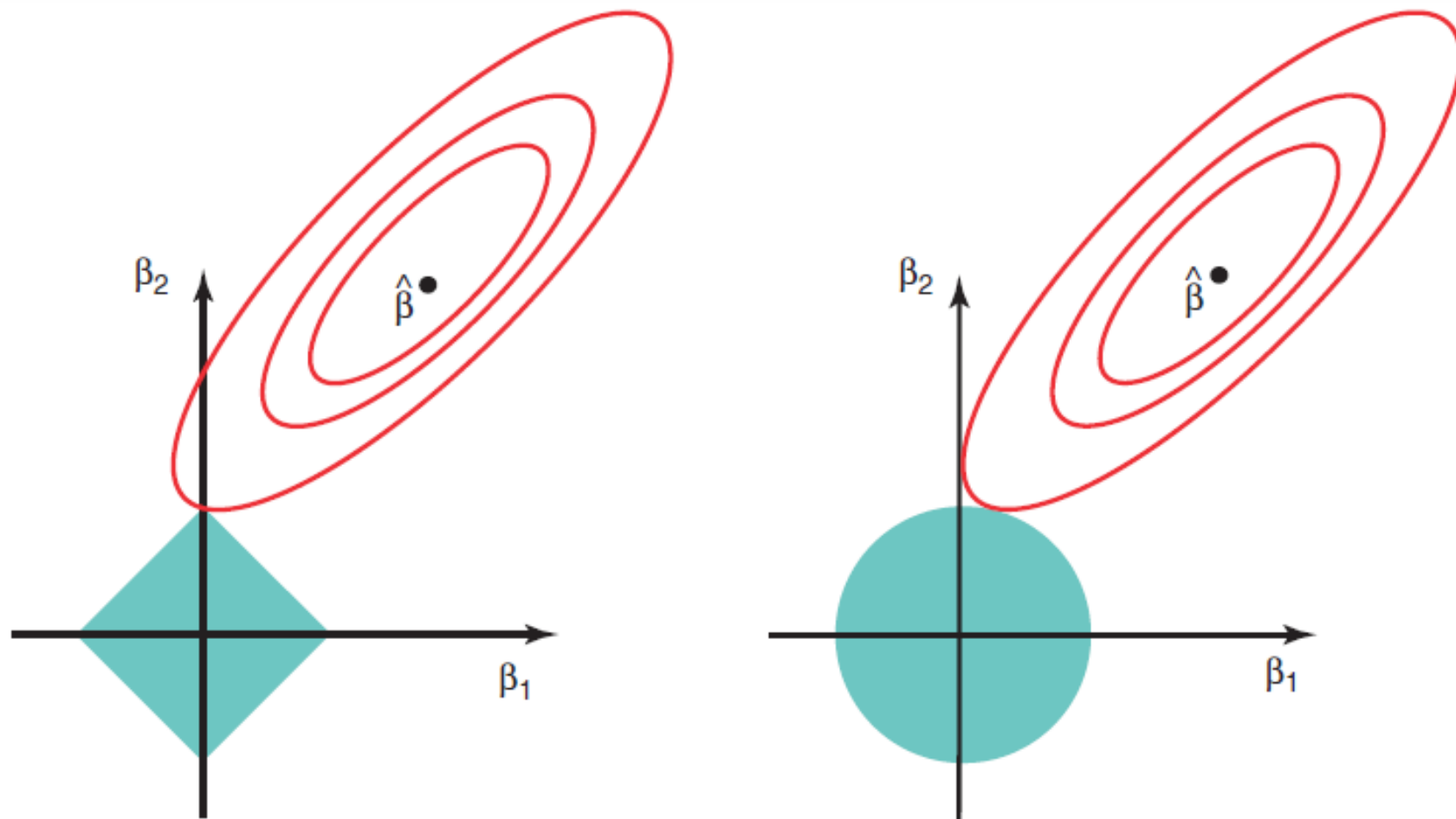


FIGURE 6.7. Contours of the error and constraint functions for the lasso (left) and ridge regression (right). The solid blue areas are the constraint regions, $|\beta_1| + |\beta_2| \leq s$ and $\beta_1^2 + \beta_2^2 \leq s$, while the red ellipses are the contours of the RSS.

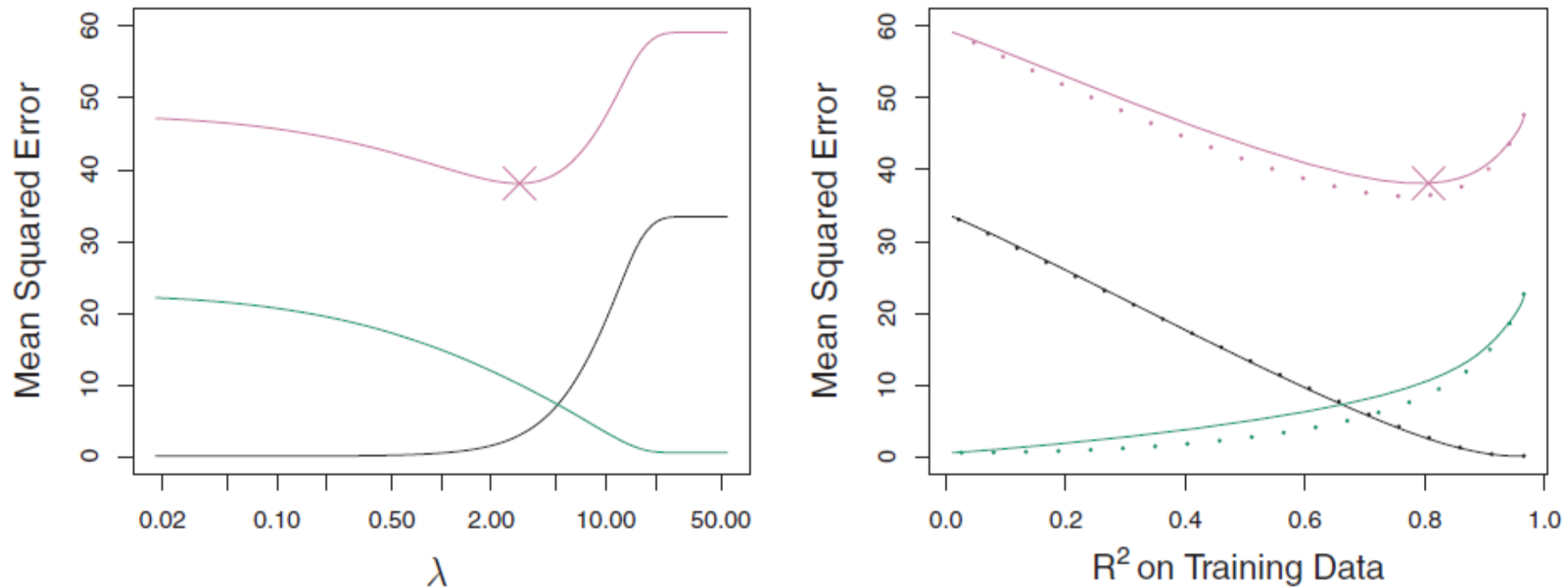


FIGURE 6.8. Left: Plots of squared bias (black), variance (green), and test MSE (purple) for the lasso on a simulated data set. Right: Comparison of squared bias, variance and test MSE between lasso (solid) and ridge (dashed). Both are plotted against their R^2 on the training data, as a common form of indexing. The crosses in both plots indicate the lasso model for which the MSE is smallest.

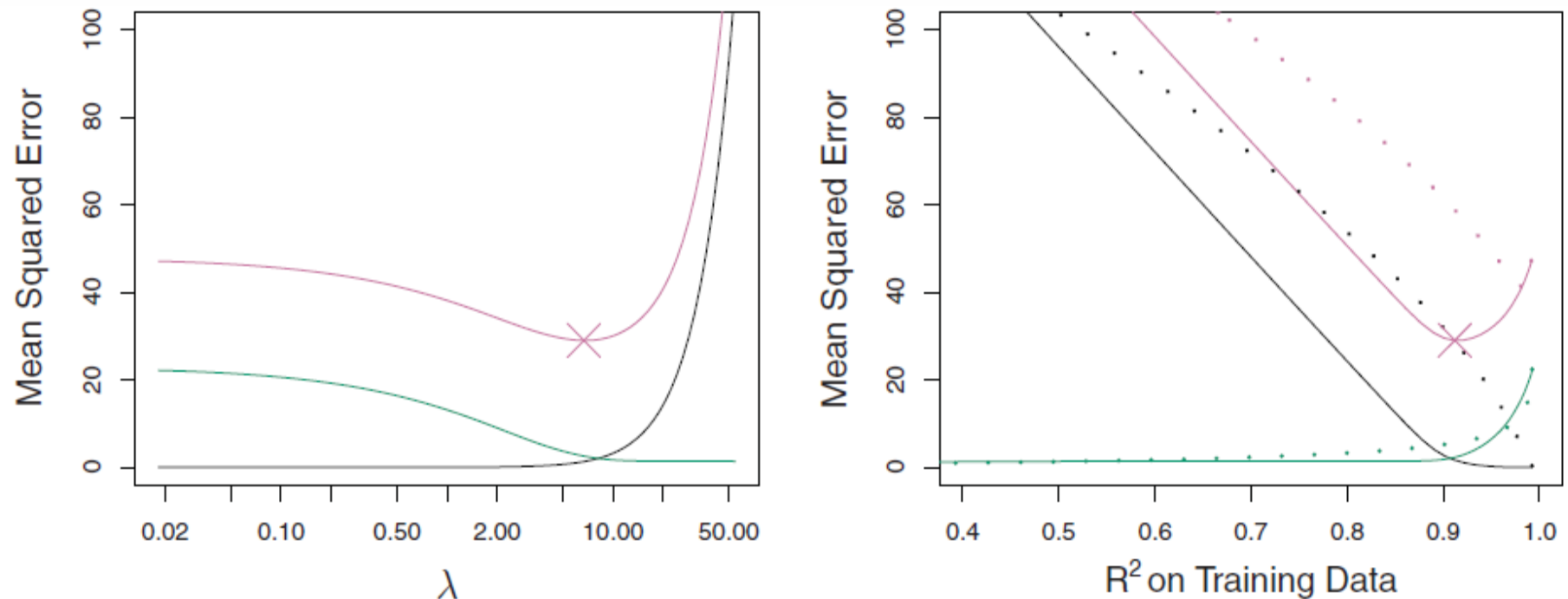


FIGURE 6.9. Left: Plots of squared bias (black), variance (green), and test MSE (purple) for the lasso. The simulated data is similar to that in Figure 6.8, except that now only two predictors are related to the response. Right: Comparison of squared bias, variance and test MSE between lasso (solid) and ridge (dashed). Both are plotted against their R^2 on the training data, as a common form of indexing. The crosses in both plots indicate the lasso model for which the MSE is smallest.

- The posterior for the coefficients can be written as

$$p(\beta|X, Y) \propto f(Y|X, \beta)p(\beta|X) = f(Y|X, \beta)p(\beta)$$

- Assuming the usual linear model

$$Y = \beta_0 + X_1\beta_1 + \dots + X_p\beta_p + \epsilon$$

- Having independent errors drawn from a normal distribution
- If we assume $p(\beta) = \prod_{j=1}^p g(\beta_j)$
 - In Ridge regression we assume a Gaussian prior with zero mean and variance being a function of lambda
 - In Lasso we assume a double-exponential (Laplace) with zero mean and scale parameter a function of lambda

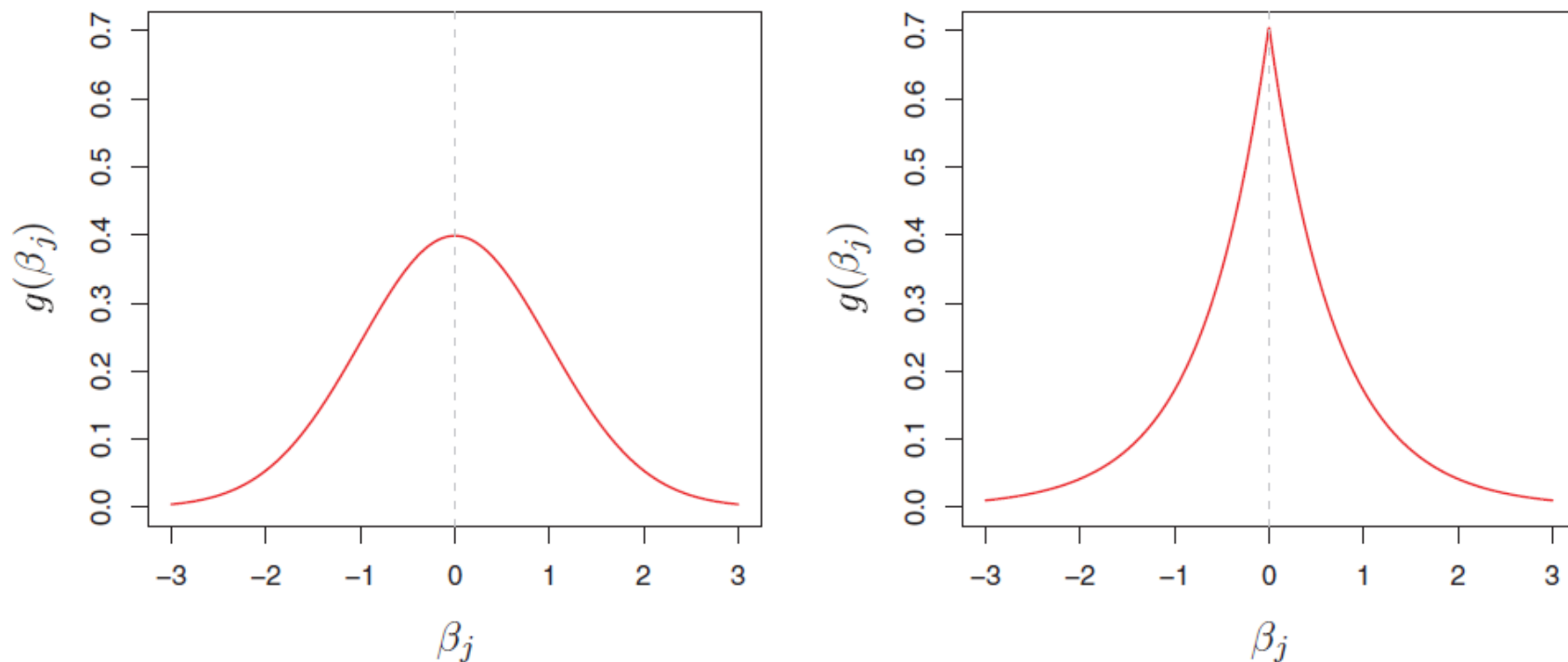


FIGURE 6.11. Left: Ridge regression is the posterior mode for β under a Gaussian prior. Right: The lasso is the posterior mode for β under a double-exponential prior.

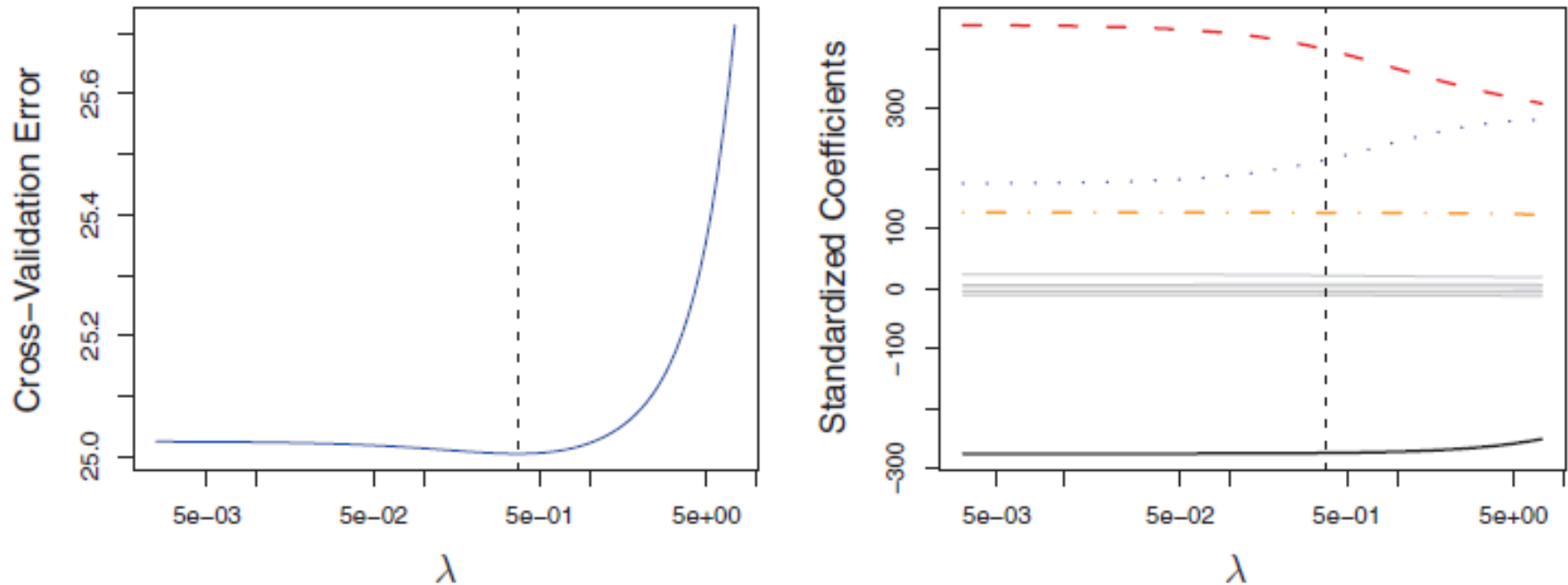


FIGURE 6.12. Left: Cross-validation errors that result from applying ridge regression to the **Credit** data set with various value of λ . Right: The coefficient estimates as a function of λ . The vertical dashed lines indicate the value of λ selected by cross-validation.

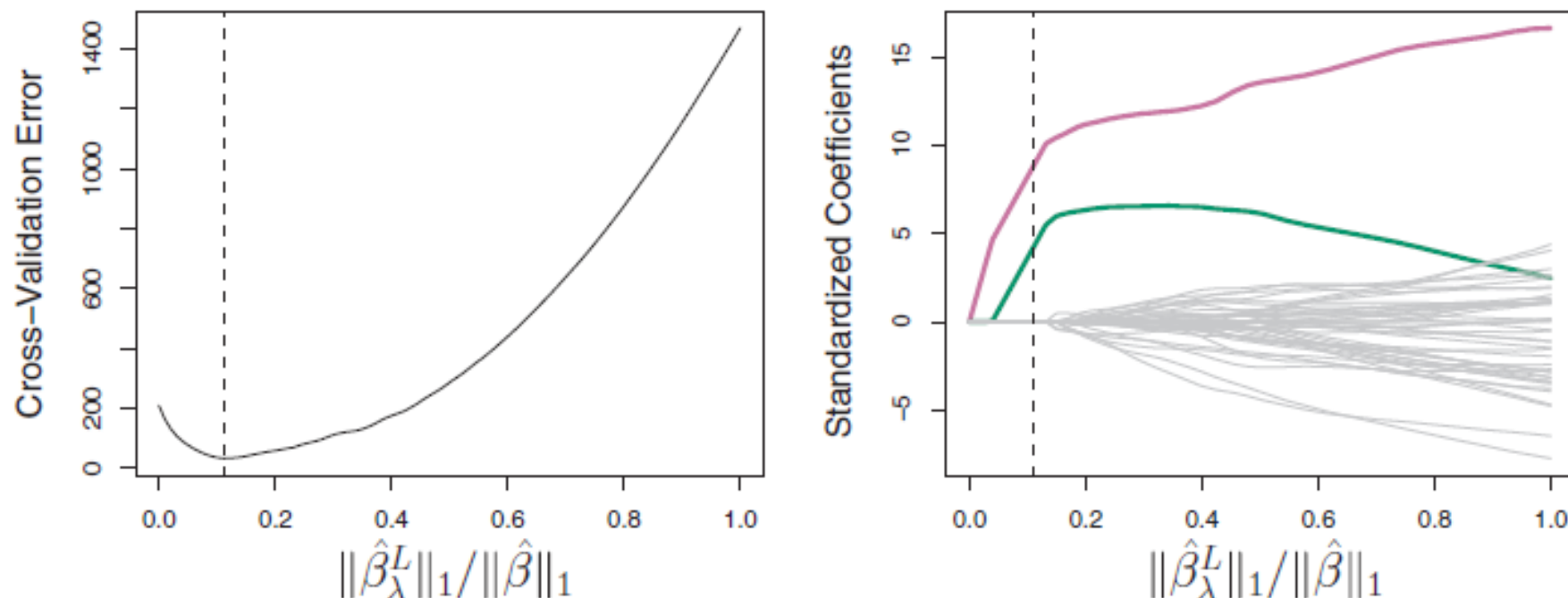


FIGURE 6.13. Left: Ten-fold cross-validation MSE for the lasso, applied to the sparse simulated data set from Figure 6.9. Right: The corresponding lasso coefficient estimates are displayed. The vertical dashed lines indicate the lasso fit for which the cross-validation error is smallest.

